

**Variable Neighborhood Search for
Finding Parameter Values in
Finite Mixture Model**

A. Bessadok, P. Hansen,
A. Rebaï

G-2009-07

February 2009

Variable Neighborhood Search for Finding Parameter Values in Finite Mixture Model

Adel Bessadok

Pierre Hansen

*GERAD and HEC Montréal
3000, chemin de la Côte-Sainte-Catherine
Montréal (Québec) Canada, H3T 2A7
{adel.bessadok, pierre.hansen}@gerad.ca*

Abdelwaheb Rebaï

*Department of Quantitative Methods
Faculté des Sciences Economique et de Gestion Sfax
B.P. 1088
Sfax 3018, Tunisia*

February 2009

Les Cahiers du GERAD

G-2009-07

Copyright © 2009 GERAD

Abstract

Finding maximum likelihood parameter values for Finite Mixture Model (FMM) is often done with the Expectation Maximization (EM) algorithm. However the choice of initial values can severely affect the time to attain convergence of the algorithm and its efficiency in finding global maxima. We alleviate this defect by embedding the EM algorithm within the variable Neighborhood Search (VNS) metaheuristic framework. Computational experiment in several problems in literature as well as some larger ones are reported.

Key Words: Maximization algorithm, Metaheuristic, Variable Neighborhood Search, Maximum Likelihood Estimation, Finite Gaussian Mixture Model, Global Optimization.

Résumé

Les techniques de classification probabilistes ont montrées des résultats très satisfaisants dans plusieurs domaines d'applications. Les modèles de mélange gaussien fini (MMGF) sont les plus répandus pour les données de grandes dimensions. La méthode de Maximum de Vrai Semblance (MVS) reconnue comme l'une des meilleures méthodes d'estimation des paramètres du MMGF se fait via l'algorithme itératif le plus utilisé, à savoir l'algorithme d'Estimation-Maximisation (EM). Cependant, le choix de la valeur initiale peut affecter sévèrement la convergence de l'algorithme EM ainsi que l'effcience d'atteindre le maximum global. Nous proposons de soulager cette faiblesse par l'incorporation de l'algorithme EM dans le cadre de la métaheuristique de Recherche à Voisinage Variable (RVV). Nous étudions l'application du nouvel algorithme proposé sur plusieurs problèmes avec divers degrés de complexités.

1 Introduction

Finite Mixture Models (FMM) are strong tools for modeling of a wide variety of random phenomena and to cluster data sets [2]. they provide an efficient platform for apprehending data with complex structure; (see [8]). Because of their usefulness as a very flexible method of modeling, FMM have proved to be of great interest over the years, both in theory and practice; (see [63]). Many problems of biology, physics and social sciences are modelled using a finite mixture of distributions; (see [55]). Recently mixture model-based methods has become very popular in cluster analysis, (see [42]) for an application of FMM to micro array data.

Modeling using mixture distributions consists in the determination of the estimation of model parameters. Maximum Likelihood Estimation (MLE) is considered to be the best method among many others, such as the method of moments used by Pearson [54] and by Cohen [12], and the graphical techniques deployed during the early and mid-1900's by Harding [31] and extended by Cassie [11] (see [21]). For mixture models problems, the likelihood equations are almost always nonlinear and can not be solved by analytic means. Consequently, one must resort to seeking an approximate solution via some iterative procedure. Our main interest here, is in a special iterative method called by Dempster et al. [14] EM algorithm.

However, the EM algorithm has some weaknesses in practice. Indeed it does not automatically provide an estimate of the covariance matrix of the parameter estimates, it is sometimes very slow to converge and in some problems, the E- or M-steps may be analytically intractable, and may converge to local optima [43, 62]. Moreover, one should emphasize that in the FMM the likelihood function is usually not unimodal and the likelihood equation has multiple roots corresponding each to local maxima; hence the EM algorithm will be very sensitive to the choice of starting values. Wu [72] reported that in general, if the log-likelihood has several (local or global) maxima and stationary points, like in the case of FMM, convergence of the EM sequence to either type of points depends on the choice of starting point.

Many variants of the original EM algorithm have been proposed in this last decade by several authors to overcome this local maxima problem. McLachlan [43] proposed the use of principle components to provide suitable starting values in FMM context. Wright and Kennedy [71] use interval analysis methods to locate multiple stationary points of a log likelihood within any designated region of the parameter space. Ueda and Nakano [66] presented a deterministic annealing EM algorithm (DAEM) method for the EM iterative process in order to be able to recover from a poor choice of starting value. By introducing the temperature parameters to modify the posteriori probability in the E-step, they provided the EM with the ability to avoid the local maxima in some specific cases but in the process make it slow so this approach may not be appropriate in general (see [45]). Moreover, the DAEM and other similar extensions of EM are useless with respect to the problem of inappropriate distribution of the components in data space when locally trapped. Recently, in order to circumvent the local optimum problem of the EM algorithm for parameter estimation in a FMM, Ueda et al. [67] proposed a split-and-merge EM (SMEM) algorithm in which they applied a split-and-merge operation to the usual EM algorithm. The basic idea of the SMEM algorithm is the following: after convergence of the usual EM algorithm, one first uses the split-and-merge operation to update the values of some parameters among all the parameters, then one performs the next round of the usual EM algorithm, and alternatively iterate the split-and-merge operation and the EM algorithm until some stopping criterion is met. Obviously, this not only benefits from the appealing simplicity of the usual EM algorithm, but also improves its global convergence capability. Vlassis and Likas [68] proposed a Greedy EM algorithm for learning a FMM to overcome the limitation of getting trapped in one of the many local maxima of the likelihood function when using the EM algorithm.

The choice of initial values is considered as a crucial point in the algorithm-based literature, as it can severely affect the time to convergence of the algorithm and its efficiency to pinpoint the global maxima (see [10]). Finch et al. [19] used a quasi-Newton method as an iterative algorithm and proposed, for two component Gaussian mixture, that only the mixing weight should be given an initial value and the rest of the parameters be automatically estimated based on this values. Karlis and Xekalki [36] presented a brief review of such method. In recent paper Biernacki, Celeux, and Govaert [5], propose a method for getting the highest

likelihood value in the framework of FMM: they identify a strategy which is based on random initialization of EM, characterized in three steps search/run/select in a fixed number of iterations. Biernacki [6] proposes a strategy to initialize the EM algorithm in FMM context by defining a starting value distribution on a mixture parameter space including all possible EM trajectories.

In this paper we develop a new method that combines a metaheuristic variable neighborhood search (VNS) algorithm with the EM algorithm to overcome as much as possible the local maxima problem.

The present work thus illustrates the association of these two algorithms to overcome the negative effect of the poor starting value choices; it is organized in six sections: (i) after this introduction (ii) we give a general presentation of the FMM and MLE in literature and formulation of the data to be treated, (iii) introduction and application of the EM algorithm for the FMM, (iv) presentation of the VNS algorithm associated with EM algorithm as a basic solution (v) some implementation issues, and simulation as an experimental results (vi) summary and discussion of the general procedure and results of the new approach in conclusions.

2 General Model and Data Presentation

In this section we shortly review and describe the mathematical framework and concepts behind FMM models followed by the MLE procedure and assumption.

2.1 Mixture Model Review

According to review of the literature, the first published investigation relating to finite mixture models dates back to the work of Newcomb [53] and Pearson. In such models, it is assumed that a sample of observations is drawn from a specified number of the underlying populations of unknown proportions. A particular form of the distributions of the observations in each of the underlying populations is specified, and the purpose of the finite mixture approach is to decompose the sample into its mixture components. Specific forms of distributions supposing a continuous variable which have been extensively used include the Gaussian ([13, 33, 69]), exponential ([64, 65]) and with a discrete variable Bernoulli distribution typically known as latent structure models ([26, 38]). However, most work on Finite Gaussian Mixture Model (FGMM) was intended either to simplify or to offer more accessible estimates in restricted cases.

2.2 Finite Mixture Model Formulation

As a simple and general presentation of mixture distributions we suppose a random variable, X , takes values in a sample space, Ω , with probability distribution $P(x)$, where x is its realization. The probability distribution can be written in the form of

$$P(x) = \sum_{j=1}^k \pi_j P_j(x) \quad (1)$$

where $P_j, j = 1, 2, \dots, k$ are the components distribution of the mixture, verifying a probability distribution proprieties. The $\pi_j, j = 1, 2, \dots, k$ are called the mixing weight where $\pi_j > 0$ for $j = 1, 2, \dots, k$ and $\sum_{j=1}^k \pi_j = 1$. Frequently, component distributions are assumed to have non parametric forms. In this case, they are parametrized by the elements of a set $\alpha = \alpha_1, \alpha_2, \dots, \alpha_k$ where α_k is the unknown vector parameters from the k^{th} component of the mixture. The Eq. (1) will be

$$P(x) = \sum_{j=1}^k \pi_j P_j(x|\alpha_j) \quad (2)$$

Thus, $\theta = \{\pi_1, \pi_2, \dots, \pi_k, \alpha_1, \alpha_2, \dots, \alpha_k\}$ is considered to be the complete collection of all FMM parameters. The mixture distribution then takes the form

$$P(x|\theta) = \sum_{j=1}^k \pi_j \tilde{P}_j(x|\alpha_j) \quad (3)$$

where, each of $\alpha_1, \alpha_2, \dots, \alpha_k$ belongs to the same parameter space, denoted by Θ and $\tilde{P}_j(\cdot|\alpha_j)$ is assumed to be a conditional distribution in each component. In this case $\pi = (\pi_1, \pi_2, \dots, \pi_k)$ may be defined as a probability distribution over θ , where $\pi_j = Pr(\alpha = \alpha_j), j = 1, 2, \dots, k$.

2.3 Maximum Likelihood Estimation Review

Our main objective is to estimate θ as mixture model parameters. From the time of the appearance of Pearson's paper [54] only the method of moments was usually the method of choice for estimating the FMM parameters. Following, the attention was focused on graphical techniques as an alternative for numerical analysis, ([7, 11, 21, 31]). With the arrival of increasingly powerful computers and increasingly sophisticated numerical methods during the 1960's, investigators began to turn to the method of MLE, originally developed by Fisher in the 1920's [20] (see [1, 59]), as the most widely preferred approach to mixture density estimation problems. Hasselblad [33] treated MLE for mixtures of any number of univariate Gaussian densities. The general case of mixture of any number of multivariate Gaussian densities was considered by Wolf [69]. Hosmer [40] compared the maximum likelihood estimates for two univariate Gaussian densities obtained from three different types of samples. Fryer and Robertson, proved that the maximum likelihood has been shown to be superior to the method of moments for the estimation of finite mixtures, from this time, the likelihood approach for finite normal mixtures has become increasingly popular.

2.4 Maximum Likelihood Estimation Formulation

Given $X = \{X_1, X_2, \dots, X_i, \dots, X_m\}$, considered as m independent observations from the mixture, their joint probability distribution is the product of each individual distribution. Therefore the likelihood function is given by

$$P(X|\theta) = \prod_{i=1}^m \left[\sum_{j=1}^k \pi_j \tilde{P}_j(x_i|\alpha_j) \right] \quad (4)$$

The maximum likelihood principle states that we should choose as an estimate of θ , a value of the observed data x which maximizes $P(x|\theta)$. That is,

$$\theta^* = \arg \max_{\theta} P(X|\theta). \quad (5)$$

θ^* is called the MLE of θ . In order to estimate θ , it is typical to introduce the log likelihood function defined as

$$L(\theta) = \ln P(X|\theta) \quad (6)$$

Since $\ln P(x|\theta)$ is a strictly increasing function, the value of θ which maximizes $P(x|\theta)$ also maximizes $L(\theta)$. MLE corresponds to a solution of the following likelihood equation:

$$\partial L / \partial \theta = 0. \quad (7)$$

Involving the log of the sum makes the maximization of L numerically difficult. Eq. (7) becomes non linear and closed form solution of this equation cannot be found. Therefore, some iterative methods should be applied.

The likelihood log function associated to this model comprises multiple local maxima. In such situations, the MLE must be sought numerically using non linear optimization algorithms and it may possible to compute iteratively the MLE by using iterative procedures. There are many general iterative procedures which are suitable for finding an approximate solution of likelihood equations. Rao [56] and Mendenhall [47] developed iterative procedures which successfully used to obtain approximate solutions of nonlinear equations satisfied by MLE. The main methods deployed are the Newton-Raphson maximization procedure or some variant such as Fisher's scoring method and quasi-Newton methods. Our main interest here, however, is in a special iterative method which is unrelated to the above ones and which has been applied to a wide variety of mixture problems over the last fifteen or so years called the EM algorithm.

3 EM Algorithm

In this section we introduce a briefly review of the literature on the EM algorithm. After this, we present an overview derivation of the EM algorithm as an alternative solution for missing or incomplete data, some proprieties are also discussed.

3.1 EM Algorithm Review

The spirit of EM algorithm was revealed and deployed independently by several different researchers, such as Newcomb [53] in his paper titled “A generalized Theory of the Combination of Observations so as to Obtain the Best Result” he figured out the two steps of the algorithm, McKendrick [41] analyzed an epidemic study linked to EM algorithm described in Meng [51], Hartley [32] had described the development the main idea of the EM algorithm and many other authors cited by Meng and Van Dyk [49] and McLachlan and Krishnan [42] until Dempster et al. [14] brought their ideas together, proved convergence, and coined the term “EM algorithm”. Since the publication of Dempster et al. [14], an abundant number of papers have been published employing the EM algorithm in many areas. Meng and Pedlow [50] found over 1000 EM-related articles appearing in almost 300 journals. Typical application areas of the EM algorithm include genetics and mixture distributions parameters estimation such as presented by Tan and Chang [60]. Further applications in many different fields and contexts can be found in [39, 42, 45, 46, 57]. Therefore this algorithm is a largely applicable approach to the iterative computation of MLE, useful in a variety of missing or incomplete-data problems.

The EM algorithm has several appealing properties: it is numerically stable, iterations are computationally attractive, the algorithm can usually be implemented easily, reliable global convergence is ensured, each EM iteration increase the likelihood function. In particular, it is generally easy to program and requires small storage space (see [42]). In fact the cost per iteration is generally low, which can offset the larger number of iterations needed for the EM algorithm compared to other competing procedures such as Newton-Raphson and it can be used to provide estimates of missing data. The EM algorithm is well known as a computationally simpler algorithm for obtaining MLE.

3.2 EM Derivation

In the comments following in Eq. (7), estimating FMM parameters using MLE amounts to solving a non-linear system of equations. However, the intuitive idea behind EM algorithm shows that, if we know the component of the mixture form which an observed data point is generated then the problem would be simpler and we get a closed-form solution of Eq. (7). Since this information is not known when the data is observed, the sample of data is then termed as incomplete data. Define the complete data to be the fully categorized data and the problem will be reformulated as an incomplete or hidden problem. Let $C = (O, H)$ being a sample of the complete data where $O = \{o_1, o_2, \dots, o_i, \dots, o_m\}$ is the sample of m observed data and $H = \{h_1, h_2, \dots, h_i, \dots, h_m\}$ defined as a sample of m hidden data. Each $h_i = (h_{i1}, h_{i2}, \dots, h_{ij}, \dots, h_{ik})$ where $h_{ij} \in \{0, 1\}$ corresponding to the mixture component to which o_i belongs. Thus, the likelihood function for the complete data can be written as

$$P(C|\theta) = \prod_{i=1}^m \prod_{j=1}^k [\pi_j P_j(o_i|\alpha_j)]^{h_{ij}} \quad (8)$$

This function is considered to be as joint complete distribution where the marginal distribution of O will be

$$P(O|\theta) = \sum_h P(C|\theta) = \sum_z P(O|h, \theta)P(h|\theta) \quad (9)$$

The goal of EM in its basic idea is to find θ such that the likelihood $P(O|\theta)$ or equivalently $L(\theta)$ is maximized. EM algorithm is an iterative procedure for maximizing $L(\theta)$. Assume that after the n^{th} iteration the current

estimate for θ is given by $\theta^{(n)}$. Since the objective is to maximize $L(\theta)$, we wish to compute an update estimate θ , such that

$$L(\theta) > L(\theta^{(n)}) \quad (10)$$

Equivalently, we want to find a new updated parameter, say $\theta^{(n+1)}$ such

$$\theta^{(n+1)} = \arg \max_{\theta} \{L(\theta) - L(\theta^{(n)})\} \quad (11)$$

then the difference is

$$\begin{aligned} L(\theta) - L(\theta^{(n)}) &= \log(\sum_h P(O|h, \theta)P(h|\theta)) - \log(P(O|\theta^{(n)})) \\ &= \log\left(\frac{\sum_h P(O|h, \theta)P(h|\theta)}{P(O|\theta^{(n)})}\right) \end{aligned} \quad (12)$$

However, using Bayes rules we get

$$P(h|O, \theta^{(n)}) = \frac{P(O, h|\theta^{(n)})}{P(O|\theta^{(n)})} \quad (13)$$

the Eq. (12) could be write with

$$L(\theta) - L(\theta^{(n)}) = \log\left[\frac{\sum_h P(O|h, \theta)P(h|\theta).P(h|O, \theta^{(n)})}{P(O, h|\theta^{(n)})}\right] \quad (14)$$

Notice that this expression involves the logarithm of a sum. Since $P(h|O, \theta^{(n)})$ is a probability measure, we have that $P(h|O, \theta^{(n)}) \geq 0$ and $\sum_h P(h|O, \theta^{(n)}) = 1$. For instance, we may apply Jensen's inequality¹ to get

$$L(\theta) - L(\theta^{(n)}) \geq \sum_h P(h|O, \theta^{(n)}) \log\left(\frac{P(O|h, \theta)P(h|\theta)}{P(O, h|\theta^{(n)})}\right) \quad (15)$$

$$\triangleq \Delta(\theta|\theta^{(n)}) \quad (16)$$

equivalently, Eq. (11) will be,

$$\theta^{(n+1)} = \arg \max_{\theta} \{\Delta(\theta|\theta^{(n)})\} \quad (17)$$

$$= \arg \max_{\theta} \sum_h P(h|O, \theta^{(n)}) \log P(O, h|\theta) \quad (18)$$

$$= \arg \max_{\theta} E_{h|O, \theta^{(n)}} \log P(O, h|\theta) \quad (19)$$

$$= \arg \max_{\theta} E_{h|O, \theta^{(n)}} \mathbb{L}(C|\theta) \quad (20)$$

$$= \arg \max_{\theta} Q(\theta; \theta^{(n)}) \quad (21)$$

In going from Eq. (17) to Eq. (18) we drop terms which are constant with respect to θ . Hence, the Eq. (20) shows well the two EM algorithm steps which are:

1. **E-step:** Compute the conditional expectation of the complete data log likelihood $L_c(\theta)$, given the observed data O , using the current iteration $\theta^{(n)}$ for θ .
2. **M-step:** Update the value of θ , say $\theta^{(n+1)}$, that maximizes the E-step. The **E-** and **M-**steps are continuously repeated until the difference $|\theta^{(n+1)} - \theta^{(n)}|$ or $|L(\theta^{(n+1)}) - L(\theta^{(n)})|$ changes by an arbitrarily small amount.

¹for constants $\lambda_i \geq 0$ with $\sum_i \lambda_i = 1$ it shown that $\log \sum_i \lambda_i x_i \geq \sum_i \lambda_i \log(x_i)$

3.3 EM Convergence

The crucial property of the EM algorithm proved by Dempster et al. (1977) is that the observed data log-likelihood $L(\theta)$ can never decrease during the EM sequence. This process continues until $L(\theta)$ converges. In fact we had in Eq. (16) that

$$\begin{aligned} L(\theta) &\geq L(\theta^{(n)}) + \Delta(\theta|\theta^{(n)}) \\ &\triangleq \ell(\theta|\theta^{(n)}) \end{aligned}$$

Additionally, it is straightforward to show that $\Delta(\theta^{(n)}|\theta^{(n)}) = 0$, hence,

$$\ell(\theta|\theta^{(n)}) = L(\theta^{(n)}) \quad (22)$$

thus, the $\ell(\theta|\theta^{(n)})$ function is bounded above by the $L(\theta)$ likelihood function. The following Figure 1, illustrates the EM procedure for two iterations.

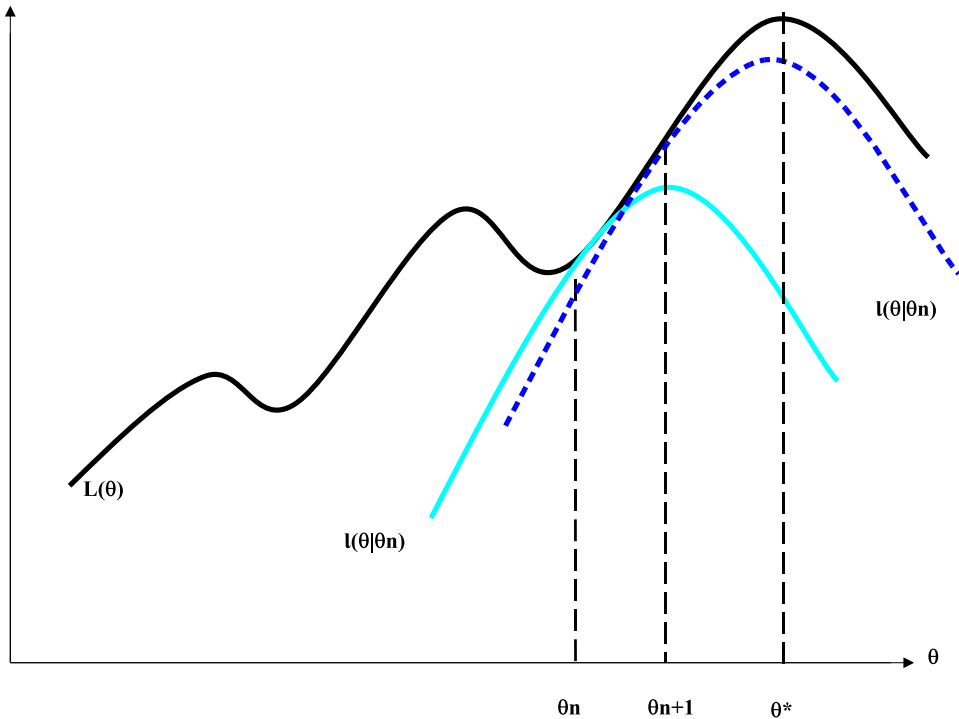


Figure 1: EM computes the function $\ell(\theta)$ using the current estimate $\theta^{(n)}$ and choose the update estimate $\theta^{(n+1)}$ as the maximum point of $\ell(\theta)$. In the next iteration at θ^* the same $\ell(\theta)$ will be generated causing the algorithm to end.

Once our objective is to find $\theta^{(n+1)}$ that maximize $L(\theta)$, we can deduce from Figure 1, that $\ell(\theta^{(n+1)}|\theta^{(n)}) \geq L(\theta^{(n)}|\theta^{(n)}) = L(\theta^{(n)})$. Therefore, at each iteration, $L(\theta)$ cannot decrease.

3.4 EM and Mixture Model

In MM context, the complete data likelihood in Eq. (8) is

$$L(C|\theta) = \sum_{i=1}^m \sum_{j=1}^k h_{ij} \log\{\pi_j P_j(o_i|\alpha_j)\} \quad (23)$$

Since $L(C|\theta)$ is a linear function of the hidden indicator variables h_{ij} , the E - step is reduces to the computation of the conditional expectation of h_{ij} , which, given an observed data o_i , using the current estimate $\theta^{(n)}$ for θ , is computed as

$$\tau_{ij}^{(n)} = E[H_{ij}|o_i; \theta^{(n)}] = P(H_{ij} = 1|o_i; \theta^{(n)}) \quad (24)$$

for each i, j , which is the current estimate of the posterior probability of the i^{th} observation generated from j component conditional on P_i and $\theta^{(n)}$ given by

$$\tau_{ij}^{(n)} = \frac{\pi_j^{(n-1)} P_j(o_i; \alpha_j^{(n-1)})}{\sum_{j=1}^k \pi_j^{(n-1)} P_j(o_i; \alpha_j^{(n-1)})} \quad (25)$$

On the **M-step** of $(n + 1)^{\text{th}}$ iteration, we update the value of θ , say $\theta^{(n+1)}$, that maximizes

$$Q(\theta; \theta^{(n)}) = \sum_{i=1}^m \sum_{j=1}^k \tau_{ij}^{(n)} \log\{\pi_j P_j(o_i|\alpha_j)\} \quad (26)$$

For MM, the estimation of mixing weight is done by differentiating

$$Q(\theta; \theta^{(n)}) - \lambda \left(\sum_{j=1}^k \pi_j - 1 \right)$$

with respect to π_j and setting derivative equal to zero, where λ is a lagrange multiplier, one has

$$\pi_j^{(n+1)} = \frac{1}{m} \sum_{i=1}^m \tau_{ij}^{(n)} \quad (27)$$

As for the updating of θ , it is obtained as an appropriate root of

$$\sum_{i=1}^m \sum_{j=1}^k \tau_{ij}^{(m)} \frac{\partial \log P_j(o_i|\alpha_j)}{\partial \theta} = 0 \quad (28)$$

In the FMM context as shown in example of Figure 1, the likelihood function will most probably have many local maxima, especially when the number of mixture-components is large (see [35]). However, performing of the EM algorithm in FMM context will provide a local maxima of the likelihood function of the observed data tend to converge to a local optima, and not necessarily the global one as results by Xu and Jordan [73]. Consequently, the EM algorithm will be very sensitive to the choice of the initial value $\theta^{(0)}$. Specifically, the effectiveness of the EM algorithm considerably depends on this first initialization. In the next section, we will present a new variant of the EM algorithm to alleviate the influence of the initial values on the performance of FMM estimation parameters.

4 Embedding EM in VNS

In this section, we propose a way to alleviate the local optima problem encountered when using EM alone. We first present an overview of the literature and describe in general terms the basic rules of the VNS metaheuristic.

4.1 EM and Global Optimization Problem

When local optima enumerates the EM algorithm can, and often does, lead to one of them instead of tow the global maxima. In fact, as shown in Section 3.2, in the M-step of the EM algorithm we are concerned with finding $\theta^{(n+1)}$ such that

$$\theta^{(n+1)} = \arg \max_{\theta} Q(\theta|\theta^{(n)}) \quad (29)$$

The $Q(\theta|\theta^{(n)})$ function is multimodal. However, we can formulate the problem of the form (29) as;

$$\begin{aligned} \max \quad & f(\theta) = Q(\theta|\theta^{(n)}) \\ \text{subject to} \quad & \theta \in \Theta \end{aligned} \tag{30}$$

where $f(\theta)$ is the objective function to be maximized and Θ is the set of feasible solutions. A solution $\theta^* \in \Theta$ is optimal if $f(\theta^*) \geq f(\theta)$, $\forall \theta \in \Theta$. In other words, we can view this problem as a global optimization one.

Woodruff and Reiners [70] considered the modeling of such sophisticated topics commonly attends to *NP-hard* problem. Such problems, particularly when involve continuous variables are often very difficult to solve. So one can either limit oneself to small instances solvable global optimization technique, or limit oneself to heuristic optimization. The later is in fact done by EM, but as we will show experimentally below embedding it in the VNS format leads to a more efficient algorithm in terms of values of solution obtained while steel keeping resolution time reasonable.

Combinatorial and global optimization algorithms are typically attracted in solving instances of problems that are believed to be hard in general. However, the use of available exact algorithms such branch-and-bound, cutting planes, decomposition, Lagrangian relaxation, column generation, and many others may not reach to solve very large instances. Moreover, Hansen and Mladenović [29], asserted that many practical instances of such problems of the form (30), arising in Operations Research and other fields, are too complex for an exact solution to be realized in conceivable time. Thus one has to endeavor to heuristics, which provide an approximate solution, or sometimes an optimal but without proof of its optimality but for the favor of a reasonable time realization. Local search is considered as one of the most used type of heuristic [28]. Local search algorithms move from an initial solution to another neighborhood solution in the space of candidate solutions by alternation of local changes, which improve at each time the objective function, until a solution deemed optimal is found or a time bound is elapsed. Although, an EM algorithm can be employed to perform local search for the problem of the form (30) [68].

4.2 VNS and Heuristics

Heuristic search procedures that aspire to find global optimal solutions to hard combinatorial optimization problems usually require some type of adjustment to overcome local optimality. In recent years, many authors extend this methodology and developed several metaheuristics algorithms for avoid being trapped in local optima with a poor value. The most relevant procedures in terms of their application to a wide variety of problems are: Tabu Search is by now a well-known metaheuristic for solving hard combinatorial optimization problems ([23, 24, 25]), Multi Start (MS) methods are devoted to the Monte Carlo random re-start in the context of nonlinear unconstrained optimization, where the method simply evaluates the objective function at randomly generated points [58], adaptive multi-start [9], simulated annealing [37], one of the most well known MS methods is the greedy adaptive search procedures (GRASP). The GRASP methodology was introduced by Feo and Resende [18] and many others have contributed to an abundant enhanced results for many combinatorial problems.

However, the performance apprehended and the sophistication of such heuristics makes it difficult to allocate with a accuracy the reasons for their effectiveness. Mladenović and Hansen [52] had examined a change of neighborhood in the search as a relatively unexplored reason and they proposed a new optimization technique called variable neighborhood search VNS.

VNS is a metaheuristic method that embeds a local search heuristic for solving combinatorial and global optimization problems. VNS performances systematically the idea of neighborhood change, both in ascendant to local maxima and in escape from the hills which contain them [30].

4.3 VNS Algorithm

Local search algorithms are widely applied to numerous hard computational problems, including problems from computer science and in particularly artificial intelligence, mathematics, operations research, engineer-

ing, and bioinformatics. Moreover, they often are building blocs for more sophisticated heuristics. A basic scheme for local search can be presented as follows

Initialization. Select a neighborhood structure N , that will be used in the search; find an initial solution x ;

Repeat the following until the stopping condition (i.e., finding a local optimum) is met:

- (a) Find the best neighbor $x' \in N(x)$ of x ;
 - (b) If x' is not better than x , stop. Otherwise, set $x = x'$ and return to (a);
-

Step of local search heuristic.

The stopping condition of this heuristic, using one neighborhood structure is satisfied as soon as a local optimum is reached. In our study the EM algorithm itself is considered as a local search structure. To improve upon the basic scheme so obtained, one can use a MS strategy that iterates for a number of times the local search from initial solution generated randomly until no further progress is made or an limit for computing time for the step is reach. However, for a considerable number of local optima, the best of those found by MS may be very far from the global optimum (Boese, Kahng and Muddu, [9]). Actually, the MS method concentrates in exploring many hills but without exploring properties of local optima so found.

VNS and contrary to other metaheuristics based on local search methods, does not follow a trajectory but explores increasingly distant neighborhoods of the current incumbent solution, and jumps from this solution to a new one if and only if an improvement has been made. Several questions about selection of neighborhood structures are in order [29]:

- (i) What properties of the neighborhoods are mandatory for the resulting scheme to be able to find a globally optimal or near-optimal solution?
- (ii) What properties of the neighborhoods will be helpful in finding a near-optimal solution?
- (iii) Should neighborhoods be nested? Otherwise how should they be ordered?
- (iv) What are desirable properties of the sizes of neighborhoods?

The basic VNS method described by Hansen and Mladenović ([27, 28, 29]), combines deterministic and stochastic changes of neighborhood. Denote with $N_k(k = 1, 2, \dots, k_{max})$ a finite set of preselected neighborhood structures, and with $N_k(x)$ the set of solutions in the k^{th} neighborhood of x . Its steps are given as:

Initialization. Select the set of neighborhood structures $N_k, k = 1, 2, \dots, k_{max}$, that will be used in the search; find an initial solution x ; choose a stopping condition;

Repeat the following until the stopping condition is met:

1. set $k \leftarrow 1$;
 2. Repeat the following steps until $k = k_{max}$:
 - (a) Shaking. Generate a point x' at random from the k^{th} neighborhood of x ($x' \in N_k(x)$);
 - (b) Local search. Apply some local search method with x' as initial solution; denote with x'' the so obtained local optimum;
 - (c) Move or not. If this local optimum is better than the incumbent, move there ($x \leftarrow x''$), and continue the search with $N_1(k \leftarrow 1)$; otherwise, set $k \leftarrow k + 1$;
-

Steps of the basic VNS

The stopping condition criteria could be such as maximum fixed number of iterations, maximum CPU time allowed, or maximum number of iterations since the last increase in the Log likelihood function. One of the major challenges in the metaheuristic VNS is the selection of neighborhood structures properties and desirable properties of the sizes of neighborhoods in away to be able to find a global optimal or best optimal

solution and to do so fairly in a reasonable time realization. In fact, To avoid being blocked in a hill, while there may be higher ones, Hansen and Mladenović [29] suggested that the union of the neighborhoods around any feasible solution θ should contain the whole feasible set:

$$\Theta \subseteq N_1(x) \cup N_2(x) \cup \dots \cup N_{k_{max}}(x), \forall x \in X.$$

These sets may cover X without necessarily partitioning it, which is easier to implement, e.g. when using nested neighborhoods, i.e.,

$$N_1(x) \subset N_2(x) \subset \dots \subset N_{k_{max}}(x), \forall x \in X.$$

If these properties do not hold, one might still be able to explore X completely, by traversing small neighborhoods around parameters values on some trajectory, but it is no more guaranteed. For instance, we define a first neighborhood $N_1(x)$ as a subdivision of the interval data range and then iterating it k times to obtain neighborhoods $N_k(x)$ for $k = 2, \dots, k_{max}$. They have the property that their sizes are increasing. Therefore if, as is often the case, one goes many times through the whole sequence of neighborhoods the first ones will be explored more thoroughly than the last ones.

4.4 EM and VNS

To let EM algorithm to be not totally depending on the first initialization is to reformulated it using the VNS method. We may consider EM as a local search in global optimization context and estimating the parameters model by maximizing the Log likelihood function subject to each parameters belong to the set of feasible solutions. We define the neighborhood structures as subintervals obtained from the data distribution range. The algorithm will be a combination of the EM and VNS (EMVNS). Therefore the basic EMVNS steps are:

Initialization. Choose an initial solution θ ; select the set of neighborhood structure by defining the intervals range I_p for the means, covariances and mixing weights parameters and by choosing the maximum number of embedded intervals in I_p (I_{pk} , $k = 1, 2, \dots, k_{max}$); choose a stopping condition, that will be used in the perturbation phase; choose a stopping condition;

Repeat the following until the stopping condition is met:

1. set $k \leftarrow 1$;
 2. Repeat the following steps until $k = k_{max}$:
 - (a) Perturbation. Generate a parameter θ' at random from the k^{th} neighborhood of θ ($\theta' \in I_{pk}(\theta)$);
 - (b) Application of the EM algorithm with θ' as initial solution; denote with θ'' the so obtained local optimum;
 - (c) Move or not. If this local optimum is better than the incumbent, move there ($\theta \leftarrow \theta''$), and continue the search with $I_{p1}(k \leftarrow 1)$; otherwise, set $k \leftarrow k + 1$;
-

Steps of the basic EMVNS

A general procedure of the EMVNS approach is presented in Figure 2.

5 Application to Finite Gaussian Mixture Model

In this section we apply our method to eight FGMM examples with different degree of complexity in order to show the effectiveness of the EMVNS approach comparing to MS method over these degree of the problem complexity.

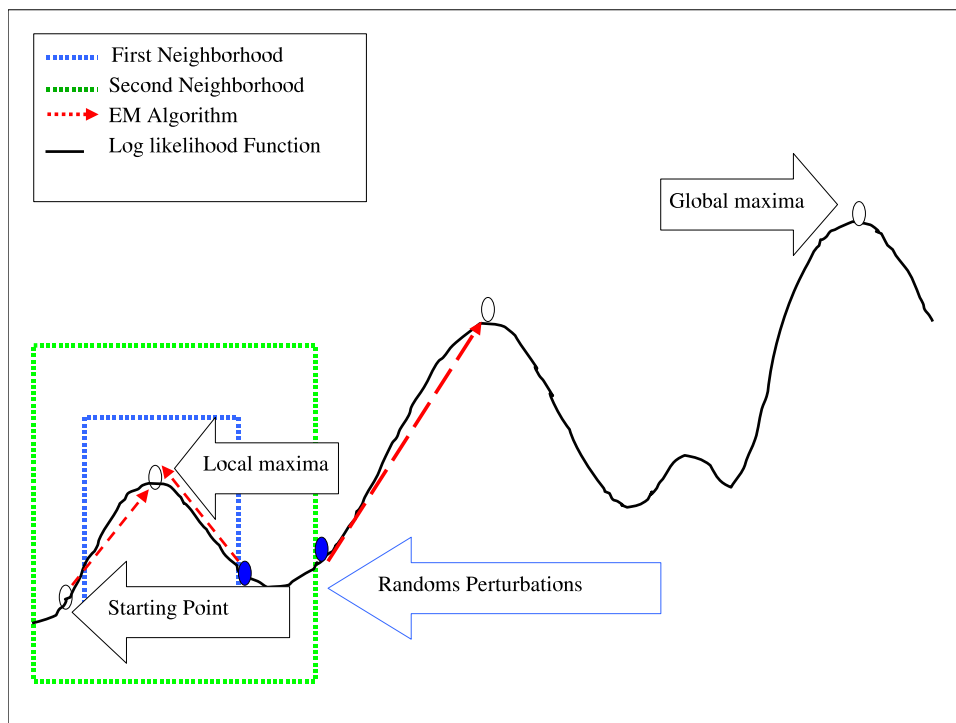


Figure 2: General procedure of EMVNS scheme

5.1 EM and FGMM

To establish FGMM parameters estimation, we explicitly derive the EM steps for Finite d-dimensional Gaussian Mixture Model. The mixed weight π_j is the unknown probability of occurrence of the j^{th} component in the mixture. Assume that each parameters Gaussian components α_j has a vector mean μ_j and covariance matrix $\Sigma_j = \sigma^2 \mathbf{I}$ where Σ_j is a positif definite symetric matrix. The marginal Finite d-dimensional Gaussian Model distribution is given by

$$P(O) = \sum_{j=1}^k \frac{\pi_j}{(2\pi)^{d/2} |\Sigma_j|^{1/2}} \exp\left\{-\frac{1}{2}(o - \mu_j)^t \Sigma_j^{-1} (o - \mu_j)\right\} \quad (31)$$

The parameters to be estimated are $\alpha_j = (\mu_j, \sigma_j)$, and π_j , $j = 1, 2, \dots, k$. Then using the two steps for estimating the model parameters denoted by, $\theta = (\mu_j, \sigma_j, \pi_j; j = 1, 2 \dots, k)$ show that E-step:

$$\tau_{ij}^{(n)} = \frac{\sigma_j^{-d} \exp\{-\|o - \mu_j(n)\|^2 / 2\sigma_j^2(n)\}}{\sum_{j=1}^k \sigma_j^{-d}(n) \exp\{-\|o - \mu_j(n)\|^2 / 2\sigma_j^2(n)\}} \quad (32)$$

M-step:

$$\pi_j^{(n+1)} = \frac{1}{m} \sum_{i=1}^m \tau_{ij} \quad (33)$$

$$\mu_j^{(n+1)} = \frac{\sum_{i=1}^m \tau_{ij} o_i}{\sum_{i=1}^m \tau_{ij}} \quad (34)$$

$$\sigma_j^{(n+1)} = \frac{\sum_{i=1}^m \tau_{ij} \|o_i - \mu_j(n+1)\|^2}{\sum_{i=1}^m \tau_{ij}} \quad (35)$$

We denote by c-G-d-MM the c component of Gaussian with d dimensional Mixture Model (e.g. 6G2MM is FGMM with 6 components in two dimension space). The eight examples chosen are displayed in Table 1.

Table 1: Examples treated

| MODEL | DATA | MEAN | COVARIANCE | MIXING WEIGHT π |
|--------|---------|--|--|--|
| 2G1MM | REAL | $\mu = [15; 9.5]$ | $\sigma = [2.1; 2.1]$ | $[0.3; 0.7]$ |
| | INITIAL | $\mu = [10.76; 16.62]$ | $\sigma = [1.6; 0.01]$ | $[0.98; 0.02]$ |
| 3G1MM | REAL | $\mu = [-5; 0; 2]$ | $\sigma = [0.5; 0.5; 0.5]$ | $[0.5; 0.3; 0.2]$ |
| | INITIAL | $\mu = [0.73; -4.72; -5.86]$ | $\sigma = [1.14; 0.23; 0.10]$ | $[0.49; 0.44; 0.07]$ |
| 2G2MM | REAL | $\mu_1 = [0; 0]$ $\mu_2 = [0; 0]$ | $\Sigma_1 = [3, 0; 0, 1/3]$ $\Sigma_2 = [1/3, 0; 0, 3]$ | $[0.7, 0.3]$ |
| | INITIAL | $\mu_1 = [0.14; -0.04]$ $\mu_2 = [-1.82; 1.94]$ | $\Sigma_1 = [1, 0.3; 0.3, 1]$ $\Sigma_2 = [1, 0.5; 0.5, 1]$ | $[0.96, 0.04]$ |
| 3G2MM | REAL | $\mu_1 = [0; 1]$ $\mu_2 = [0; -1]$ $\mu_3 = [-1; 2]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $= [0.2, 0.1; 0.1, 0.2]$ | $[0.3; 0.3; 0.4]$ |
| | INITIAL | $\mu_1 = [-1.53; 1.73]$ $\mu_2 = [-0.08; 1.92]$ $\mu_3 = [0.52; 0.19]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $= [1, 0.3; 0.3, 1]$ | $[0.46; 0.01; 0.53]$ |
| 4G2MM | REAL | $\mu_1 = [0; -2]$ $\mu_2 = [2, 0]$ $\mu_3 = [0; 2]$ $\mu_4 = [-2, 0]$ | $\Sigma_1 = \Sigma_3 = [3, 0; 0, 1/3]$ $\Sigma_2 = \sigma_4 = [1/3, 0; 0, 3]$ | $[0.25, 0.25, 0.25, 0.25]$ |
| | INITIAL | $\mu_1 = [-0.3; 1.1]$ $\mu_2 = [1.9; 0]$ $\mu_3 = [-1.6; 1.2]$ $\mu_4 = [0.6; 0.1]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $= \Sigma_4 = [0.6, 0.9; 0.9, 0.4]$ | $[0.41, 0.26, 0.18, 0.16]$ |
| 6G2MM | REAL | $\mu_1 = [0.75; -0.5]$ $\mu_2 = [0.5; 1]$ $\mu_3 = [0; 1.5]$ $\mu_4 = [-1; -0.5]$ $\mu_5 = [-1.5; 0]$ $\mu_6 = [1; -1.5]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $= [0.05, 0; 0, 0.2]$ | $[1/6; 1/6; 1/6; 1/6; 1/6; 1/6]$ |
| | INITIAL | $\mu_1 = [-0.3; 1.1]$ $\mu_2 = [1.9; 0]$ $\mu_3 = [-1.6; 1.2]$ $\mu_4 = [0.6; 0.1]$ $\mu_5 = [0; 1.5]$ $\mu_6 = [1.1; -0.5]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $= [0.6, 0.9; 0.9, 0.4]$ | $[0.11; 0.20; 0.18; 0.16; 0.17; 0.18]$ |
| 8G2MM | REAL | $\mu_1 = [1.5; 0]$ $\mu_2 = [1; 1]$ $\mu_3 = [0; 1.5]$ $\mu_4 = [-1; 1]$ $\mu_5 = [-1.5; 0]$ $\mu_6 = [-1; -1]$ $\mu_7 = [0; -1.5]$ $\mu_8 = [1; -1]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $\Sigma_7 = \Sigma_8$ $= [0.01, 0; 0, 0.1]$ | $[1/8; 1/8; 1/8; 1/8; 1/8; 1/8; 1/8; 1/8]$ |
| | INITIAL | $\mu_1 = [-0.3; 1.4]$ $\mu_2 = [1.9; -0.2]$ $\mu_3 = [-1.6; 1.2]$ $\mu_4 = [0.6; 0.1]$ $\mu_5 = [0.3; -0.7]$ $\mu_6 = [1.11; -0.5]$ $\mu_7 = [1.6; -0.5]$ $\mu_8 = [1.2; 1.3]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $\Sigma_7 = \Sigma_8$ $= [0.06, 0.95; 0.95, 0.48]$ | $[0.01; 0.23; 0.29; 0.12; 0.04; 0.06; 0.17; 0.08]$ |
| 10G2MM | REAL | $\mu_1 = [1.25; 0]$ $\mu_2 = [1; 1]$ $\mu_3 = [0; 1.5]$ $\mu_4 = [-1; 1]$ $\mu_5 = [-1.5; 0]$ $\mu_6 = [-1.5; -1]$ $\mu_7 = [0; -1.5]$ $\mu_8 = [1; -1]$ $\mu_9 = [0.5; -1.5]$ $\mu_{10} = [1; -1.5]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $\Sigma_7 = \Sigma_8 = \Sigma_9$ $= \Sigma_{10} = [0.01, 0; 0, 0.1]$ | $[0.1; 0.1; 0.1; 0.15; 0.1; 0.15; 0.1; 0.1; 0.05; 0.05]$ |
| | INITIAL | $\mu_1 = [-0.3; 1.4]$ $\mu_2 = [1.9; -0.2]$ $\mu_3 = [-1.6; 1.2]$ $\mu_4 = [0.6; 0.1]$ $\mu_5 = [0.3; -0.7]$ $\mu_6 = [1.11; -0.5]$ $\mu_7 = [1.6; -0.5]$ $\mu_8 = [1.2; 1.3]$ $\mu_9 = [1.1; -0.15]$ $\mu_{10} = [1.5; -1.3]$ | $\Sigma_1 = \Sigma_2 = \Sigma_3$ $\Sigma_4 = \Sigma_5 = \Sigma_6$ $\Sigma_7 = \Sigma_8 = \Sigma_9$ $= \Sigma_{10} = [0.06, 0.95; 0.95, 0.48]$ | $[0.1; 0.1; 0.1; 0.15; 0.1; 0.15; 0.1; 0.1; 0.05; 0.05]$ |

5.2 Experimental Procedure

Before detailing the experimental procedure, we describe the sensibility of EM using a “poor” initialization to get a local maxima. We use 150 observations generated from 3G1MM described in Table 1. Our algorithm is applied using the same ‘poor’ starting value and as shown in Figure 3 the EMVNS can easily improve the parameters model estimation.

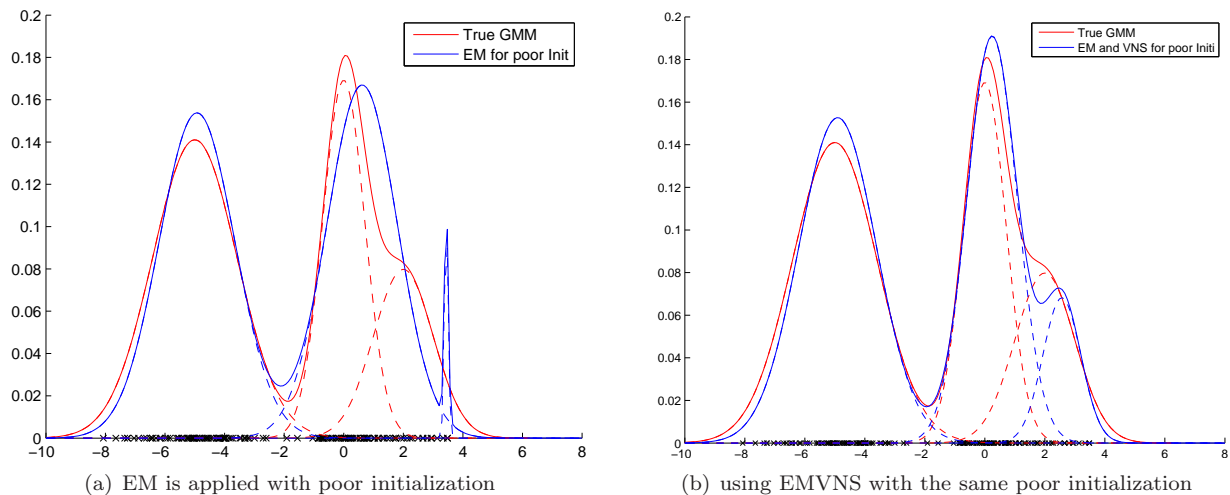


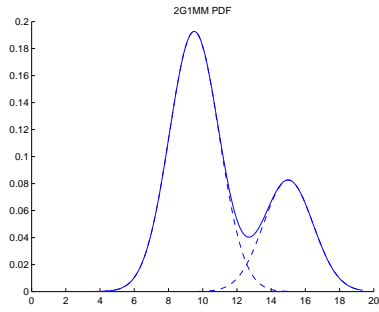
Figure 3: The use of EM can lead in general to a local maxima and not to the global one. However the implementation of EMVNS using the same poor initialization guarantee a better result.

5.2.1 Methods Deployed

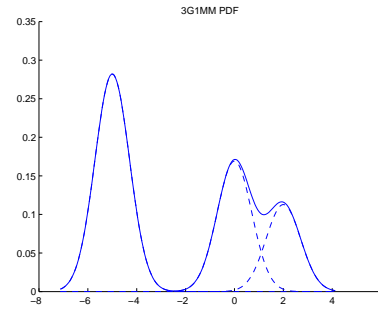
In this paper, we limit our study to comparing the EMVNS with the MS method. In deed, the MS method is the most used means of initiating EM and is considered as reference method for almost any comparison method (see [5]). In MS method the means are generated from an interval from the data range, the covariances are generated from the an interval ranging from zero to the value of sample covariances and the mixing weight are generated from a Dirichlet distribution. In EMVNS approach, we choose I_p for the means as the interval data range, for the covariances as an interval ranging from 0 to the value of sample covariances and an interval ranging from 0 to 0.9 for mixing weights. Note that EM was used in its standard form and without any acceleration scheme. To avoid the unbounded likelihoods problems, all components are chosen to have a common variance or equal determinants (see [45]).

5.2.2 Examples Treated

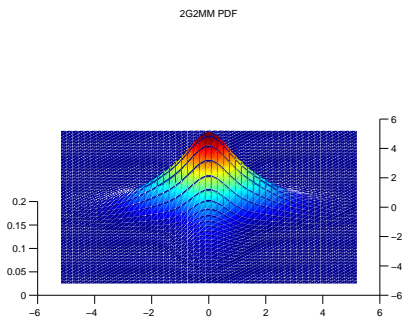
In order to give an acceptable credibility to the results provided by both methods, we choose a variety of situations, but far to be exclusive. In fact, as we will discuss later, the EM performance results are directly related to the attraction basin size and dimension of local and global maxima of the likelihood function (see [4]), thus, we considered eight examples displayed in Table 1 with different degree of complexity resumed in both component dimension and data distribution. From the Probability Density Function (PDF) of 8 FGMM examples shown in Figure 4, we can observe the diversity of data distribution from well separated as 8G2MM example to poorly separated as 2G2MM example. Figure 5 provides the information about the approximate local maxima number from a very reduced as in 2G1MM to very considerable as 10G2MM. The dendrograms in Figure 5 illustrate that the number of local maxima values is dependent on the problem complexity characterized by the number of component and data distribution.



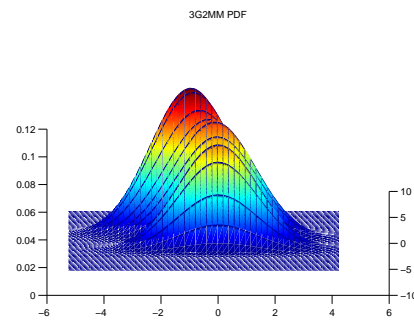
(a) A very simple FGMM



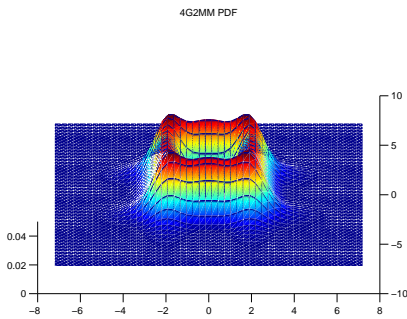
(b) FGMM in one dimension space



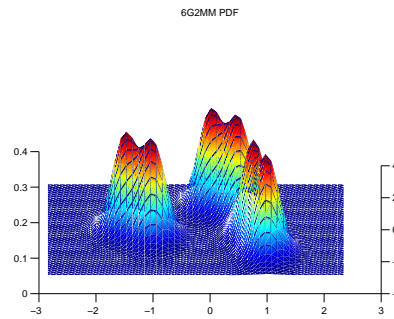
(c) The two model components are poorly separated



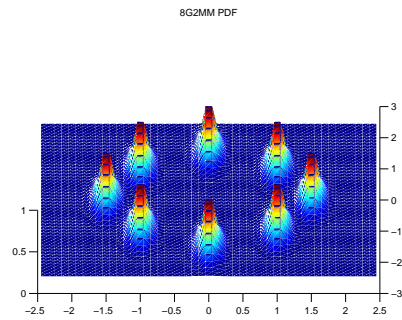
(d) The model components are relatively poor separated



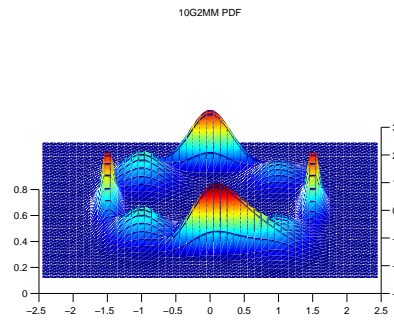
(e) The four model components are well separated



(f) The model components are quite poorly separated

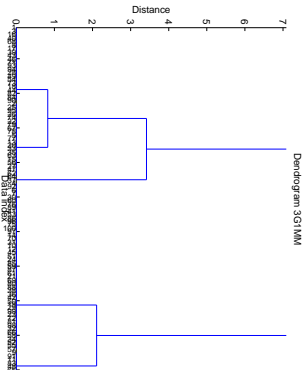
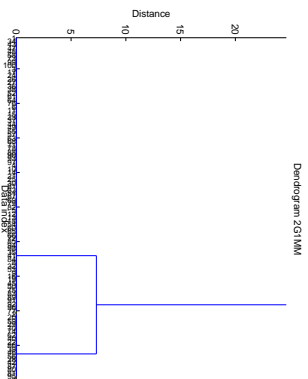


(g) The model components are very clearly separated



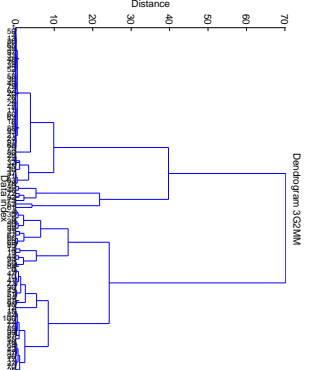
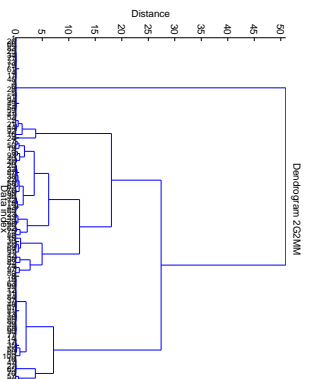
(h) The model components are quite separated

Figure 4: The PDF of eight FGMM examples are very diversified with different degree of complexity resumed in both component dimension and data distribution.



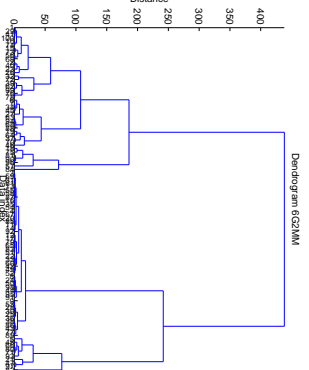
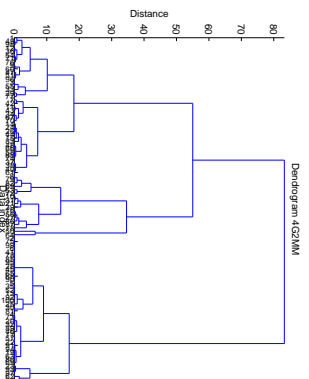
(a) The number of distinct local maxima is enormously reduced

(b) The average number of local maxima is around 3



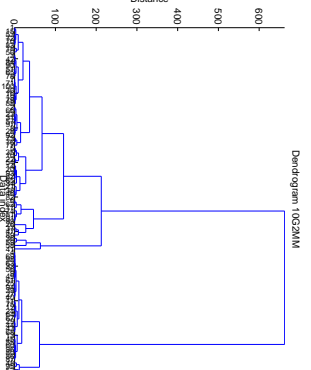
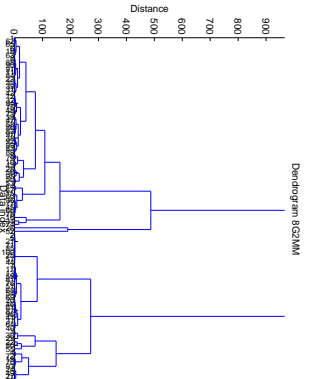
(c) The average number of local maxima is around 40

(d) The average number of local maxima is around 60



(e) The average number of local maxima is around 60

(f) The average number of local maxima is around 90



(g) The average number of local maxima is around 80

(h) The average number of local maxima is around 100

Figure 5: The dendrograms are generated from 100 iterations of the EM algorithm using the MS method as initial value and the single linkage as criteria.

5.2.3 Criteria Selected

To analyze the performance of each method we define some criteria. Our main objective is to reach the highest likelihood regardless of the time of realization even though the CPU time realization is less for EMVNS than for MS method. For instance, for the last 10G2MM example, one iteration with MS took 0.26s and only 0.14s with EMVNS. Therefore we choose a fixed number of iteration as stopping condition for both methods. To perform the competition between the two methods we limit to 100 the number of iterations relatively to the number of local maxima as presented in Figure 5. For being more realistic we choose a “poor” and random starting parameter value for EMVNS method as shown in Table 1. We fixed to 25 the maximum number of embedded intervals and for simplicity we choose the same incremental step for getting all I_{pk} as 30% of I_p (i.e. $I_{p1} = 30\%I_p$, $I_{p2} = 60\%I_p$, etc). We devote 10 trials twice for each example sample. For both methods, we record after each trial the highest Log likelihood considered as the best associated global maxima. The second criteria which is the local maxima range that give us an idea about the ability of EMVNS method to improve the search for getting the best local maxima by jumping from hill to hill and from MS method it informs us about the approximately wide range of the local maxima. The last criteria is the percentage of getting the associated global one; it characterizes the degree of complexity of the problem we treat. This percentage provided by EMVNS interpreted the percentage of time being in this global maxima hill by this method. Nevertheless, from MS method this percentage explained essentially the attraction basin size of the local maxima.

5.3 Results Obtained

The results for both poor and random and for both trials displayed in Tables 2, 3, 4, 5, 6, 7, 8, and 9 illustrated that for 2G1MM, 2G2MM and 4G2MM examples, both methods reached the same global maxima. Thus, we can depict from Figure 5 that the number of local maxima associated to 2G1MM example is very much reduced and regardless of the attraction basin size of global maxima it’s easy in this case for any simple method to gain in 100 iterations the global one. For the 2G2MM and 4G2MM took from Biernacki, Celeux and Govaert [5] examples with a considerable number of local maxima as shown in Figure 5, the percentage in getting the global maxima with MS method is very important leading to large attraction basin of global max. In this case it’s easy too to get to the global maxima regardless the components are well or poorly separated and it’s confirmed by Biernacki, Celeux and Govaert [5] results. In 8G2MM example, where the components of the model are well separated, the MS method reached in the common trials to the same global maxima as EMVNS method but as shown by Tables 5, 9, 3 and 7 it succeeded in a few trials to get the highest log likelihood. This result could be clarified by the considerable number of local maxima of the Log likelihood function with visibly large attraction basin. Thus, it is easy for MS to get the global max with a comparatively small percentage such as less than 12% and that let EMVNS in a few times passing his 100 iterations in other large attraction basin of the local maxima. Since the limit number of iterations, this is the only case where MS can perform well rather than EMVNS. For the 3G1MM considered as a simple example in one dimension space with a reduced local maxima number, the EMVNS thrived in the majority of trials to get to the highest Log likelihood. In fact, when the attraction basin size is large for a local maxima and a small for the global maxima, as argued by the quite great percentage for getting the global maxima with MS, EMVNS succeeded to jump to other local maxima until get to best hill; nonetheless, MS get trapped in this local maxima. In 6G2MM and 10G2MM examples, having a significant important number of local maxima, the competition between both methods is very rude. In deed, the percentage of getting the global maxima in MS is very small as it did not exceeded 2%, consequently, the most of local maxima had a very small attraction basin. In this case, as showing in Table 3, EMVNS had a more chance to get to the best result. The 3G2MM example resumed almost all situations discussed in the rest of the other examples. In fact, the percentage for reaching the global maxima with MS is varied from small to relatively great percentage, hence, in this situation the attraction basin size of the local maxima varied also from small to large size. In this case, EMVNS succeeded as well to get to the highest hill in the majority of trials. Therefore, we can deduct from the local maxima average and from the comparatively great percentage of getting the global maxima

with EMVNS method, that in almost all situations there is no difficulty for the EMVNS method to improve the accuracy of model parameters estimation and attain the best global maxima.

5.4 Discussions

It is well shown the impact of starting value in the degree of accuracy of FMM parameters estimation using EM algorithm. Indeed, it is clear that EM is a very sensitive to the initial values.

From the experiment results, we can briefly resume all possible situations in three cases. First, when the attraction basin of the local maxima is relatively large supported by the percentage of getting the global maxima with MS method such as more than 15%, the EMVNS guarantees to get the highest Log likelihood or at least the same than MS but not in any case the MS will have a best result than EMVNS. The second case manifested by a very important number of local maxima with relatively the same attraction basin size explained by the relatively percentage of global maxima with MS method such as between 15 to 5% and that lead, in the majority of trials, to have the same global maxima for both methods with a slight amelioration with MS method. The last case is the challenge one, because the number of local maxima is very important with a small attraction basin size argued by the small percentage in getting the global maxima with MS method that did not exceed 5% and even in this case EMVNS succeeded in the best part of trials to reach the greatest global maxima.

In fact, the large number of local maxima that conduct to large number of parameters estimation solutions, depending on the complexity of the problem. As shown, when the number of local maxima is very much reduced or when the basin attraction of the global maxima is visibly large, we can use an ordinary method as MS or any other simple method for getting best results to overcome initial values problem caused by EM algorithm. However, in practice, the data dimension is very large modeled by FMM and without guarantees to have a large basin attraction of the global max. In spite of this, the need of applying a robust method for a complex problems is necessary. As described in Section 4.2, contrarily to other methods based in local search, VNS provide a powerful and simple tool to implement for getting a best results comparing to competitive methods (see [28]). Moreover, the use of the appropriate structures in VNS leads not only to improve the Maximization of the likelihood function but with best time realization.

6 Conclusions

The choice of initial values is considered as crucial point in the algorithm-based literature as it can severely affect the time realization of convergence of the algorithm and its efficiency to pinpoint the global maxima.

A novel EMVNS algorithm for estimating FMM parameters is proposed in this paper to overcome one of the main drawbacks of EM algorithm often getting trapped at local maxima. The VNS method largely deployed in many examples had shown his efficient in getting best improvement results which exploits systematically the idea of neighborhood change, both in ascendant to local maxima and in escape from the hills which contain them. The algorithm is computationally efficient and easily implemented.

The experimental results of employing FGMM for a variety of degree of complexity of data dimension show that our algorithm can find excellent solutions with best time realization than MS method, especially in complicated situations.

The EMVNS algorithm use the VNS in his basic scheme and is focus on the estimation of FMM parameters supposing the number of FMM components are known before. therefore, developing EMVNS using several VNS extensions and finding more appropriate structures for resolving such problem appears to be desirable.

Table 2: First trials of EMVNS Performance with poor starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|-------|--------|--------|------------------|------------------|----------------------|---------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 2G1MM | 100 | 1 | -235.8753 | -235.8753 | 10.1791 | 8.4856 | 98 | 94 |
| | | 2 | -234.4101 | -234.4101 | 13.2288 | 11.1921 | 96 | 95 |
| | | 3 | -239.2080 | -239.2080 | 10.4160 | 10.4160 | 97 | 92 |
| | | 4 | -238.0905 | -238.0905 | 14.4562 | 15.3911 | 99 | 94 |
| | | 5 | -235.1474 | -235.1474 | 16.2545 | 13.2589 | 94 | 96 |
| | | 6 | -236.2357 | -236.2357 | 11.6521 | 13.2587 | 91 | 93 |
| | | 7 | -233.9541 | -231.9541 | 14.6523 | 14.3251 | 93 | 96 |
| | | 8 | -237.0245 | -237.0245 | 16.2124 | 11.3547 | 98 | 94 |
| | | 9 | -238.1544 | -238.1544 | 13.2574 | 11.2541 | 92 | 93 |
| | | 10 | -238.8563 | -238.8563 | 11.6523 | 12.6523 | 96 | 91 |
| 3G1MM | 150 | 1 | -290.0475 | -286.6524 | 1.9924 | 3.3951 | 17 | 17 |
| | | 2 | -307.6195 | -307.6195 | 4.0851 | 3.4112 | 47 | 79 |
| | | 3 | -296.8935 | -296.8481 | 2.0226 | 2.0680 | 19 | 70 |
| | | 4 | -318.4914 | -316.7653 | 1.8197 | 3.5458 | 65 | 79 |
| | | 5 | -299.0124 | -299.0124 | 12.2900 | 11.4197 | 42 | 88 |
| | | 6 | -306.7865 | -304.4735 | 2.1404 | 5.0500 | 72 | 7 |
| | | 7 | -292.1598 | -292.1598 | 5.6254 | 5.2654 | 76 | 75 |
| | | 8 | -303.1527 | -303.0161 | 7.0874 | 7.1822 | 6 | 68 |
| | | 9 | -312.3613 | -312.3613 | 3.1955 | 3.1675 | 74 | 91 |
| | | 10 | -301.1689 | -301.1689 | 3.1113 | 3.1113 | 65 | 74 |
| 2G2MM | 200 | 1 | -641.2487 | -641.2487 | 46.5814 | 22.3507 | 36 | 76 |
| | | 2 | -632.5786 | -632.5786 | 45.3939 | 3.4112 | 27 | 79 |
| | | 3 | -600.5044 | -600.5044 | 62.6585 | 48.2360 | 40 | 93 |
| | | 4 | -620.1937 | -620.1937 | 36.1978 | 34.3310 | 31 | 82 |
| | | 5 | -627.4071 | -627.4071 | 53.7468 | 42.9049 | 17 | 18 |
| | | 6 | -649.4448 | -649.4448 | 57.4785 | 56.8515 | 29 | 86 |
| | | 7 | -611.8998 | -611.8998 | 47.3719 | 32.4887 | 37 | 96 |
| | | 8 | -614.0520 | -614.0520 | 24.3866 | 17.9981 | 3 | 80 |
| | | 9 | -601.6633 | -601.6633 | 36.8120 | 34.5052 | 34 | 96 |
| | | 10 | -639.6007 | -639.6007 | 51.6233 | 32.7748 | 12 | 73 |
| 3G2MM | 200 | 1 | -652.5217 | -648.1404 | 47.2736 | 14.9758 | 6 | 83 |
| | | 2 | -646.0257 | -622.5508 | 57.1086 | 23.4749 | 2 | 49 |
| | | 3 | -605.1128 | -584.6639 | 65.7443 | 20.4489 | 35 | 33 |
| | | 4 | -655.8673 | -655.8673 | 53.0580 | 3.7038 | 17 | 87 |
| | | 5 | -610.5412 | -612.5774 | 71.2584 | 12.5263 | 12 | 20 |
| | | 6 | -630.8122 | -630.8122 | 54.6093 | 6.9880 | 16 | 81 |
| | | 7 | -614.8514 | -614.8514 | 61.9125 | 21.7185 | 14 | 68 |
| | | 8 | -643.9878 | -639.1087 | 64.8320 | 6.9985 | 1 | 64 |
| | | 9 | -653.0959 | -652.4204 | 67.4756 | 2.8292 | 1 | 34 |
| | | 10 | -646.1953 | -646.1953 | 66.0131 | 4.2911 | 29 | 93 |

Table 3: First trials of EMVNS Performance with poor starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|--------|--------|--------|--------------------|--------------------|----------------------|------------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 4G2MM | 200 | 1 | -770.1656 | -770.1656 | 79.2570 | 57.8955 | 26 | 90 |
| | | 2 | -749.1386 | -749.1386 | 89.8699 | 78.8091 | 24 | 74 |
| | | 3 | -741.2514 | -741.2514 | 74.5417 | 65.2155 | 21 | 71 |
| | | 4 | -760.7066 | -760.7066 | 68.8641 | 61.2430 | 32 | 55 |
| | | 5 | -776.4492 | -776.4492 | 81.8411 | 59.6209 | 40 | 79 |
| | | 6 | -744.2546 | -744.2546 | 80.3412 | 59.7544 | 23 | 81 |
| | | 7 | -751.1382 | -751.1382 | 76.6545 | 66.2541 | 24 | 80 |
| | | 8 | -748.2112 | -748.2112 | 81.6995 | 72.4121 | 23 | 73 |
| | | 9 | -752.2243 | -752.2243 | 82.2546 | 77.5546 | 25 | 78 |
| | | 10 | -761.9663 | -761.9663 | 72.8910 | 66.2096 | 37 | 66 |
| 6G2MM | 500 | 1 | -1.0041e+03 | -999.7815 | 421.4680 | 349.8224 | 1 | 1 |
| | | 2 | -958.0554 | -958.0554 | 450.9760 | 379.5808 | 2 | 64 |
| | | 3 | -989.4930 | -987.6576 | 423.2859 | 357.9227 | 1 | 20 |
| | | 4 | -1.0137e+03 | -1.0137e+03 | 383.7455 | 4328.2597 | 1 | 42 |
| | | 5 | -953.4334 | -944.4864 | 468.9808 | 407.8856 | 1 | 54 |
| | | 6 | -989.7588 | -981.1017 | 428.3935 | 354.6687 | 1 | 23 |
| | | 7 | -985.7730 | -969.1041 | 439.6149 | 370.9343 | 1 | 5 |
| | | 8 | -1.0255e+03 | -1.0255e+03 | 429.1275 | 330.0983 | 1 | 64 |
| | | 9 | -1.0173e+03 | -1.0083e+03 | 394.4401 | 350.0143 | 1 | 77 |
| | | 10 | -1.0300e+03 | -1.0314e+03 | 398.5337 | 315.0695 | 1 | 45 |
| 8G2MM | 700 | 1 | -1.0042e+03 | -1.0042e+03 | 1.0154e+03 | 967.0880 | 8 | 55 |
| | | 2 | -1.0437e+03 | -1.0437e+03 | 938.1917 | 910.6738 | 4 | 53 |
| | | 3 | -973.2614 | -973.2614 | 1.0736e+03 | 974.4371 | 8 | 76 |
| | | 4 | -961.2759 | -961.2759 | 1.0453e+03 | 989.5428 | 6 | 52 |
| | | 5 | -1.0055e+03 | -1.0055e+03 | 1.0059e+03 | 967.4489 | 5 | 56 |
| | | 6 | -978.1461 | -978.1461 | 1.0165e+03 | 982.0156 | 12 | 41 |
| | | 7 | -973.6314 | -1.2444e+03 | 1.0098e+03 | 587.4345 | 12 | 27 |
| | | 8 | -1.0117e+03 | -1.0117e+03 | 1.0023e+03 | 954.3250 | 8 | 18 |
| | | 9 | -1.0513e+03 | -1.1150e+03 | 988.7539 | 917.9848 | 8 | 54 |
| | | 10 | -974.6313 | -974.6313 | 1.0133e+03 | 923.3240 | 10 | 85 |
| 10G2MM | 800 | 1 | -1.1078e+03 | -1.1022e+03 | 1.2953e+03 | 1.0856e+03 | 1 | 66 |
| | | 2 | -1.1656e+03 | -1.1623e+03 | 1.2360e+03 | 1.0111e+03 | 1 | 10 |
| | | 3 | -1.1161e+03 | -1.1109e+03 | 1.2805e+03 | 1.0489e+03 | 1 | 72 |
| | | 4 | -1.1512e+03 | -1.1392e+03 | 1.2298e+03 | 1.0308e+03 | 1 | 29 |
| | | 5 | -1.1552e+03 | -1.1552e+03 | 1.2319e+03 | 954.9397 | 1 | 42 |
| | | 6 | -1.1330e+03 | -1.1681e+03 | 1.1882e+03 | 1.0101e+03 | 1 | 16 |
| | | 7 | -1.1360e+03 | -1.1341e+03 | 1.2245e+03 | 1.0442e+03 | 1 | 55 |
| | | 8 | -1.1137e+03 | -1.1850e+03 | 1.2506e+03 | 1.0221e+03 | 1 | 1 |
| | | 9 | -1.1642e+03 | -1.1560e+03 | 1.2532e+03 | 1.0519e+03 | 1 | 48 |
| | | 10 | -1.1731e+03 | -1.1684e+03 | 1.2141e+03 | 1.0132e+03 | 1 | 37 |

Table 4: First trials of EMVNS Performance with random starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|-------|--------|--------|------------------|------------------|----------------------|-------------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 2G1MM | 100 | 1 | -234.8521 | -234.8521 | 11.1269 | 9.5632 | 94 | 91 |
| | | 2 | -236.5411 | -236.5411 | 16.2541 | 3.1695 | 95 | 98 |
| | | 3 | -239.6056 | -239.6056 | 13.8891 | 1.9895e-13 | 96 | 100 |
| | | 4 | -236.2693 | -236.2693 | 13.6254 | 7.3254 | 95 | 97 |
| | | 5 | -228.3542 | -228.3542 | 4.5214e-13 | 2.9562e-13 | 100 | 100 |
| | | 6 | -237.1642 | -237.1642 | 14.2671 | 10.1547 | 91 | 93 |
| | | 7 | -226.8521 | -226.8521 | 12.0652 | 4.0516e-13 | 92 | 100 |
| | | 8 | -235.1245 | -235.1245 | 16.2547 | 15.2541 | 96 | 92 |
| | | 9 | -234.0251 | -234.0251 | 12.3658 | 10.3596 | 91 | 94 |
| | | 10 | -220.8229 | -220.8229 | 17.3641 | 3.9790e-13 | 99 | 100 |
| 3G1MM | 150 | 1 | -290.1672 | -290.1672 | 55.7924 | 0.2672 | 35 | 96 |
| | | 2 | -298.0510 | -296.8191 | 2.7856 | 1.2320 | 74 | 25 |
| | | 3 | -288.9858 | -288.9070 | 2.0077 | 1.5669 | 8 | 71 |
| | | 4 | -293.7079 | -293.7079 | 1.0692 | 1.0692 | 64 | 99 |
| | | 5 | -281.0147 | -281.0147 | 4.7916 | 1.3642e-12 | 72 | 100 |
| | | 6 | -289.9147 | -289.5405 | 2.9777 | 2.0934 | 77 | 60 |
| | | 7 | -276.1808 | -276.1808 | 12.4557 | 11.2750 | 70 | 98 |
| | | 8 | -307.4659 | -306.6627 | 2.4843 | 3.2876 | 70 | 84 |
| | | 9 | -296.5679 | -295.2190 | 4.7682 | 1.3489 | 81 | 71 |
| | | 10 | -304.4261 | -304.4261 | 6.9570 | 6.2528e-13 | 74 | 100 |
| 2G2MM | 200 | 1 | -643.3488 | -643.3488 | 48.9028 | 1.2506e-12 | 38 | 100 |
| | | 2 | -618.4387 | -618.4387 | 39.2941 | 18.7931 | 16 | 98 |
| | | 3 | -629.2587 | -629.2587 | 62.3592 | 44.0413 | 23 | 98 |
| | | 4 | -627.5673 | -627.5673 | 37.3981 | 3.4106e-13 | 30 | 100 |
| | | 5 | -581.3516 | -581.3516 | 12.2900 | 1.1369e-12 | 40 | 98 |
| | | 6 | -631.5586 | -631.5586 | 43.3517 | 3.1243 | 29 | 81 |
| | | 7 | -601.5012 | -601.5012 | 57.6534 | 39.2371 | 41 | 92 |
| | | 8 | -625.3583 | -625.3583 | 41.2587 | 3.8471e-13 | 31 | 100 |
| | | 9 | -585.2847 | -585.2847 | 22.12544 | 1.21459e-12 | 40 | 100 |
| | | 5 | -635.7149 | -635.7149 | 38.8518 | 20.5429 | 1 | 61 |
| 3G2MM | 200 | 1 | -648.9412 | -649.4180 | 76.1761 | 2.7151 | 1 | 15 |
| | | 2 | -659.7583 | -658.1108 | 53.9242 | 6.8212e-13 | 2 | 100 |
| | | 3 | -635.7000 | -635.3556 | 51.5436 | 43.8824 | 17 | 15 |
| | | 4 | -653.9328 | -652.8712 | 59.6598 | 3.8281 | 8 | 11 |
| | | 5 | -620.8220 | -619.0844 | 55.3170 | 1.7377 | 22 | 97 |
| | | 6 | -643.2921 | -643.2921 | 66.4684 | 55.2612 | 2 | 72 |
| | | 7 | -642.6907 | -641.2610 | 63.3844 | 1.4297 | 1 | 58 |
| | | 8 | -613.5389 | -613.5389 | 51.6619 | 47.2346 | 1 | 33 |
| | | 9 | -619.8786 | -617.6769 | 65.2008 | 56.4982 | 31 | 30 |
| | | 10 | -611.3266 | -624.4769 | 77.3488 | 2.7285e-12 | 1 | 100 |

Table 5: First trials of EMVNS Performance with random starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|--------|--------|--------|--------------------|--------------------|----------------------|------------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 4G2MM | 200 | 1 | -752.1700 | -752.1700 | 82.1010 | 15.1649 | 35 | 99 |
| | | 2 | -797.6400 | -797.6400 | 67.9817 | 12.9877 | 4 | 79 |
| | | 3 | -742.1413 | -742.1413 | 79.3266 | 68.4991 | 30 | 64 |
| | | 4 | -760.7066 | -760.7066 | 68.8641 | 61.2430 | 32 | 55 |
| | | 5 | -758.0660 | -758.0660 | 66.2370 | 34.3501 | 16 | 55 |
| | | 6 | -751.1227 | -751.1227 | 87.0018 | 71.5188 | 24 | 99 |
| | | 7 | -772.2980 | -772.2980 | 64.9957 | 0.4575 | 20 | 99 |
| | | 8 | -756.3823 | -756.3823 | 79.1321 | 72.6345 | 14 | 57 |
| | | 9 | -740.3214 | -740.3214 | 84.8219 | 58.9910 | 4 | 39 |
| | | 10 | -755.9385 | -755.9385 | 87.8174 | 6.7298 | 21 | 92 |
| 6G2MM | 500 | 1 | -1.0039e+03 | -1.0002e+03 | 417.7638 | 8.8476 | 1 | 8 |
| | | 2 | -1.0264e+03 | -1.0278e+03 | 392.8282 | 247.8544 | 4 | 25 |
| | | 3 | -1.0264e+03 | -1.0056e+03 | 392.8282 | 24.5343 | 4 | 93 |
| | | 4 | -1.0072e+03 | -1.0101e+03 | 431.1072 | 333.6097 | 4 | 57 |
| | | 5 | -1.0072e+03 | -1.0068e+03 | 431.1072 | 5.8514 | 1 | 5 |
| | | 6 | -1.0172e+03 | -992.2265 | 378.7475 | 30.4581 | 1 | 42 |
| | | 7 | -1.0277e+03 | -1.0223e+03 | 387.7640 | 273.8456 | 1 | 43 |
| | | 8 | -1.0183e+03 | -1.0179e+03 | 393.6705 | 14.5584 | 3 | 8 |
| | | 9 | -1.0159e+03 | -1.0159e+03 | 400.5788 | 19.9525 | 1 | 67 |
| | | 10 | -999.3922 | -977.3866 | 450.8094 | 26.9714 | 1 | 87 |
| 8G2MM | 700 | 1 | -984.4266 | -984.4266 | 1.0154e+03 | 76.1694 | 8 | 35 |
| | | 2 | -985.7864 | -985.7864 | 1.0218e+03 | 913.5765 | 12 | 87 |
| | | 3 | -983.4621 | -983.4621 | 1.0136e+03 | 53.6341 | 4 | 84 |
| | | 4 | -968.0664 | -968.0664 | 1.0133e+03 | 967.1979 | 7 | 62 |
| | | 5 | -993.8567 | -1.0562e+03 | 995.0358 | 906.2064 | 10 | 82 |
| | | 6 | -1.0237e+03 | -1.0237e+03 | 996.4803 | 77.6997 | 7 | 99 |
| | | 7 | -1.0026e+03 | -1.0026e+03 | 981.5858 | 81.8980 | 11 | 92 |
| | | 8 | -998.8686 | -998.8686 | 1.0001e+03 | 66.4924 | 7 | 16 |
| | | 9 | -988.1429 | -988.1429 | 975.6521 | 53.8557 | 8 | 99 |
| | | 10 | -1.0147e+03 | -1.0147e+03 | 979.8081e+03 | 931.5988 | 10 | 64 |
| 10G2MM | 800 | 1 | -1.1385e+03 | -1.1366e+03 | 1.2299e+03 | 22.2510 | 1 | 38 |
| | | 2 | -1.1551e+03 | -1.1478e+03 | 1.2290e+03 | 38.0575 | 1 | 93 |
| | | 3 | -1.1093e+03 | -1.1068e+03 | 1.2574e+03 | 1.1802e+03 | 1 | 44 |
| | | 4 | -1.1694e+03 | -1.1732e+03 | 1.2309e+03 | 1.0507e+03 | 1 | 78 |
| | | 5 | -1.1483e+03 | -1.1472e+03 | 1.2040e+03 | 33.5633 | 1 | 12 |
| | | 6 | -1.1082e+03 | -1.1030e+03 | 1.3031e+03 | 10.8805 | 1 | 59 |
| | | 7 | -1.1360e+03 | -1.1341e+03 | 1.2245e+03 | 1.0442e+03 | 1 | 55 |
| | | 8 | -1.1254e+03 | -1.1154e+03 | 1.2621e+03 | 521.2814 | 1 | 12 |
| | | 9 | -1.1621e+03 | -1.1698e+03 | 1.2486e+03 | 38.1673 | 1 | 10 |
| | | 10 | -1.5471e+03 | -1.5455e+03 | 717.9538 | 684.8795 | 1 | 31 |

Table 6: Second trials of EMVNS Performance with poor starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|-------|--------|--------|------------------|------------------|----------------------|---------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 2G1MM | 100 | 1 | -235.8753 | -235.8753 | 10.1791 | 8.4856 | 98 | 94 |
| | | 2 | -236.5277 | -236.5277 | 11.5142 | 9.6211 | 91 | 96 |
| | | 3 | -238.2114 | -238.2114 | 12.8411 | 11.6244 | 94 | 96 |
| | | 4 | -234.1433 | -234.1433 | 11.3266 | 10.6521 | 94 | 98 |
| | | 5 | -238.2510 | -238.2510 | 13.5413 | 11.3210 | 91 | 93 |
| | | 6 | -235.0125 | -235.0125 | 16.3214 | 10.1048 | 98 | 94 |
| | | 7 | -236.2301 | -236.2301 | 14.0477 | 13.5109 | 91 | 96 |
| | | 8 | -237.4523 | -237.4523 | 11.2036 | 9.3254 | 92 | 96 |
| | | 9 | -235.3580 | -235.3580 | 13.0911 | 12.9154 | 92 | 91 |
| | | 10 | -233.2866 | -233.2866 | 16.8205 | 10.2144 | 90 | 92 |
| 3G1MM | 150 | 1 | -288.5110 | -288.5100 | 3.2262 | 1.4254 | 48 | 99 |
| | | 2 | -282.1450 | -278.7498 | 5.1305 | 5.8677 | 4 | 47 |
| | | 3 | -294.5143 | -294.5143 | 3.2514 | 2.6347 | 18 | 74 |
| | | 4 | -290.1672 | -290.1672 | 4.7367 | 3.2314 | 16 | 81 |
| | | 5 | -295.4229 | -290.4305 | 2.2856 | 6.3440 | 32 | 44 |
| | | 6 | -298.0510 | -296.8191 | 2.7856 | 4.01764 | 70 | 1 |
| | | 7 | -275.8961 | -275.8961 | 4.3314 | 3.8048 | 1 | 76 |
| | | 8 | -290.9747 | -289.2354 | 5.3697 | 5.3159 | 58 | 83 |
| | | 9 | -317.6993 | -317.5027 | 5.8401 | 5.3555 | 80 | 53 |
| | | 10 | -293.7200 | 292.3509 | 2.7478 | 3.8355 | 13 | 43 |
| 2G2MM | 200 | 1 | -641.2487 | -641.2487 | 46.5814 | 22.3507 | 36 | 76 |
| | | 2 | -639.6214 | -639.6214 | 41.2147 | 28.6211 | 24 | 75 |
| | | 3 | -614.6207 | -614.6207 | 51.6503 | 46.2217 | 32 | 76 |
| | | 4 | -649.0452 | -649.0452 | 41.9605 | 32.6072 | 38 | 90 |
| | | 5 | -601.2019 | -601.2019 | 48.6263 | 42.2851 | 24 | 28 |
| | | 6 | -639.3088 | -639.3088 | 48.3644 | 30.6249 | 32 | 92 |
| | | 7 | -620.3321 | -620.3321 | 51.3017 | 38.6328 | 28 | 84 |
| | | 8 | -628.3622 | -628.3622 | 50.9423 | 31.3301 | 38 | 94 |
| | | 9 | -638.0752 | -638.0752 | 42.3285 | 31.0866 | 28 | 86 |
| | | 10 | -608.6429 | -608.6429 | 45.6275 | 40.9004 | 32 | 76 |
| 3G2MM | 200 | 1 | -649.2168 | -649.2168 | 59.0776 | 50.7962 | 2 | 73 |
| | | 2 | -623.7199 | -617.6543 | 58.4495 | 62.3994 | 26 | 76 |
| | | 3 | -651.1126 | -643.0419 | 54.814 | 46.1338 | 1 | 76 |
| | | 4 | -633.3075 | -633.3075 | 59.0551 | 50.5428 | 13 | 81 |
| | | 5 | -642.5767 | -642.5767 | 58.1452 | 50.2351 | 13 | 43 |
| | | 6 | -630.5347 | -625.4970 | 54.6974 | 48.3538 | 3 | 54 |
| | | 7 | -640.8798 | -625.6880 | 59.4298 | 71.9445 | 1 | 50 |
| | | 8 | -638.4954 | -638.4954 | 53.5249 | 46.9954 | 14 | 61 |
| | | 9 | -620.5147 | -623.3256 | 61.2547 | 51.2156 | 13 | 36 |
| | | 10 | -622.9948 | -607.1670 | 62.7560 | 67.7480 | 3 | 67 |

Table 7: Second trials of EMVNS Performance with poor starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|--------|--------|--------|--------------------|--------------------|----------------------|------------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 4G2MM | 200 | 1 | -743.1162 | -743.1162 | 75.1187 | 73.6494 | 39 | 94 |
| | | 2 | -749.2199 | -749.2199 | 86.3220 | 71.0821 | 28 | 94 |
| | | 3 | -756.3678 | -756.3678 | 79.3128 | 65.2155 | 29 | 71 |
| | | 4 | -760.3703 | -760.3703 | 91.1277 | 78.7588 | 38 | 92 |
| | | 5 | -763.3643 | -763.3643 | 62.8596 | 71.2364 | 16 | 76 |
| | | 6 | -741.3652 | -741.3652 | 79.6966 | 70.9122 | 24 | 75 |
| | | 7 | -752.9941 | -752.9941 | 81.2247 | 71.3221 | 31 | 80 |
| | | 8 | -752.2411 | -752.2411 | 74.3248 | 70.6912 | 30 | 76 |
| | | 9 | -748.2610 | -748.2610 | 65.2544 | 69.6523 | 34 | 81 |
| | | 10 | -758.6311 | -758.6311 | 80.9817 | 71.9122 | 31 | 91 |
| 6G2MM | 500 | 1 | -1.0277e+03 | -1.0331e+03 | 366.3842 | 349.8224 | 1 | 40 |
| | | 2 | -977.9569 | -944.0497 | 405.4793 | 367.8062 | 3 | 24 |
| | | 3 | -1.0408e+03 | -1.0408e+03 | 399.9315 | 334.7967 | 1 | 21 |
| | | 4 | -975.4794 | -974.8529 | 429.9173 | 306.8197 | 1 | 79 |
| | | 5 | -998.3713 | -995.8678 | 423.5845 | 243.0707 | 1 | 41 |
| | | 6 | -989.1698 | -987.4594 | 417.2788 | 232.7691 | 1 | 34 |
| | | 7 | -995.4418 | -990.2266 | 434.6577 | 318.9361 | 1 | 21 |
| | | 8 | -1.0046e+03 | -1.0046e+03 | 446.2635 | 376.4621 | 1 | 58 |
| | | 9 | -1.0188e+03 | -1.0189e+03 | 429.0988 | 328.1266 | 1 | 83 |
| | | 10 | -1.0463e+03 | -1.0437e+03 | 401.8871 | 267.3516 | 1 | 9 |
| 8G2MM | 700 | 1 | 958.6579 | 958.6579 | 1.0557e+03 | 1.0242e+03 | 2 | 17 |
| | | 2 | -1.0081e+03 | -1.0081e+03 | 9967.5115 | 905.1947 | 10 | 7 |
| | | 3 | -1.0472e+03 | -1.0472e+03 | 954.5972 | 936.1669 | 5 | 28 |
| | | 4 | -1.0077e+03 | -1.0077e+03 | 972.6653 | 901.8672 | 6 | 21 |
| | | 5 | -1.0040e+03 | -1.0040e+03 | 986.8095 | 948.3994 | 11 | 72 |
| | | 6 | -1.0090e+03 | -1.0090e+03 | 983.9204 | 953.1839 | 5 | 88 |
| | | 7 | -1.0084e+03 | -1.0084e+03 | 994.5039 | 858.3365 | 6 | 27 |
| | | 8 | -984.7680 | -1.0491e+03 | 1.06044 | 908.0154 | 7 | 45 |
| | | 9 | -1.0171e+03 | -1.0941e+03 | 990.3082 | 834.4988 | 7 | 30 |
| | | 10 | -978.5214 | -978.5214 | 985.9514 | 854.2145 | 11 | 64 |
| 10G2MM | 800 | 1 | -1.1093e+03 | -1.1068e+03 | 1.2574e+03 | 1.1802e+03 | 1 | 44 |
| | | 2 | -1.1698e+03 | -1.1582e+03 | 1.2375e+03 | 1.1437e+03 | 1 | 18 |
| | | 3 | -1.1093e+03 | -1.1068e+03 | 1.2805e+03 | 1.1802e+03 | 1 | 44 |
| | | 4 | -1.1848e+03 | -1.1797e+03 | 1.2298e+03 | 1.0362e+03 | 1 | 87 |
| | | 5 | -1.1469e+03 | -1.1447e+03 | 1.2061e+03 | 1.0662e+03 | 1 | 1 |
| | | 6 | -1.1682e+03 | -1.1826e+03 | 1.2315e+03 | 1.1525e+03 | 1 | 4 |
| | | 7 | -1.1405e+03 | -1.1741e+03 | 1.2375e+03 | 1.1044e+03 | 1 | 18 |
| | | 8 | -1.1202e+03 | -1.1226e+03 | 1.2888e+03 | 1.1900e+03 | 1 | 27 |
| | | 9 | -1.1308e+03 | -1.1303e+03 | 1.2363e+03 | 1.1382e+03 | 1 | 26 |
| | | 10 | -1.0882e+03 | -1.0879e+03 | 1.3139e+03 | 1.0151e+03 | 1 | 48 |

Table 8: Second trials of EMVNS Performance with random starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|-------|--------|--------|------------------|------------------|----------------------|------------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 2G1MM | 100 | 1 | -235.8753 | -235.8753 | 10.1791 | 1.6422e-13 | 98 | 100 |
| | | 2 | -236.5277 | -236.5277 | 11.5142 | 8.2141 | 91 | 95 |
| | | 3 | -238.2114 | -238.2114 | 12.8411 | 10.3249 | 94 | 91 |
| | | 4 | -234.1433 | -234.1433 | 11.3266 | 1.5211e-13 | 94 | 100 |
| | | 5 | -238.2510 | -238.2510 | 13.5413 | 2.3521 | 91 | 98 |
| | | 6 | -235.0125 | -235.0125 | 16.3214 | 10.5211 | 98 | 96 |
| | | 7 | -236.2301 | -236.2301 | 14.0477 | 1.4136e-13 | 91 | 100 |
| | | 8 | -237.4523 | -237.4523 | 11.2036 | 8.2183 | 92 | 94 |
| | | 9 | -235.3580 | -235.3580 | 13.0911 | 6.6271 | 92 | 96 |
| | | 10 | -233.2866 | -233.2866 | 16.8205 | 11.9521 | 90 | 94 |
| 3G1MM | 150 | 1 | -288.5110 | -288.1290 | 3.2262 | 1.8074 | 48 | 41 |
| | | 2 | -282.1450 | -279.2184 | 5.1305 | 3.0781 | 4 | 18 |
| | | 3 | -294.5143 | -294.5143 | 3.2514 | 1.6270 | 18 | 86 |
| | | 4 | -290.1672 | -290.1672 | 4.7367 | 0.2672 | 16 | 88 |
| | | 5 | -295.4229 | -293.0061 | 2.2856 | 2.4169 | 32 | 84 |
| | | 6 | -298.0510 | -296.8191 | 2.7856 | 1.4551 | 70 | 35 |
| | | 7 | -275.8961 | -275.8961 | 4.3314 | 0.2888 | 1 | 99 |
| | | 8 | -290.9747 | -289.2354 | 5.3697 | 1.7393 | 58 | 52 |
| | | 9 | -317.6993 | -317.6993 | 5.8401 | 5.1589 | 80 | 87 |
| | | 10 | -293.7200 | -291.5225 | 2.7478 | 3.4538 | 13 | 71 |
| 2G2MM | 200 | 1 | -641.2487 | -641.2487 | 46.5814 | 8.5470 | 36 | 84 |
| | | 2 | -639.6214 | -639.6214 | 41.2147 | 10.9624 | 24 | 90 |
| | | 3 | -614.6207 | -614.6207 | 51.6503 | 6.4063 | 32 | 86 |
| | | 4 | -649.0452 | -649.0452 | 41.9605 | 44.9120 | 38 | 74 |
| | | 5 | -601.2019 | -601.2019 | 48.6263 | 13.5209 | 24 | 92 |
| | | 6 | -639.3088 | -639.3088 | 48.3644 | 36.0411 | 32 | 82 |
| | | 7 | -620.3321 | -620.3321 | 51.3017 | 8.6157 | 28 | 92 |
| | | 8 | -628.3622 | -628.3622 | 50.9423 | 40.6184 | 38 | 76 |
| | | 9 | -638.0752 | -638.0752 | 42.3285 | 28.5713 | 28 | 74 |
| | | 10 | -608.6429 | -608.6429 | 45.6275 | 36.2843 | 32 | 84 |
| 3G2MM | 200 | 1 | -649.2168 | -649.2168 | 59.0776 | 5.7646 | 2 | 92 |
| | | 2 | -623.7199 | -623.7199 | 58.4495 | 48.0424 | 26 | 16 |
| | | 3 | -651.1126 | -650.5690 | 54.814 | 16.0921 | 1 | 47 |
| | | 4 | -633.3075 | -633.3075 | 59.0551 | 32.7674 | 13 | 63 |
| | | 5 | -642.5767 | -642.5767 | 58.1452 | 24.5217 | 13 | 68 |
| | | 6 | -630.5347 | -625.4970 | 54.6974 | 8.2842 | 3 | 62 |
| | | 7 | -640.8798 | -634.9935 | 59.4298 | 9.8502 | 1 | 97 |
| | | 8 | -638.4954 | -638.4954 | 53.5249 | 37.5486 | 14 | 53 |
| | | 9 | -620.5147 | -621.5142 | 61.2547 | 48.2547 | 13 | 41 |
| | | 10 | 622.9948 | -621.5530 | 62.7560 | 2.7384 | 3 | 33 |

Table 9: Second trials of EMVNS Performance with random starting point

| Model | Sample | TRIALS | Global Maxima | | Local Maxima AVERAGE | | % Global Maxima | |
|--------|--------|--------|--------------------|--------------------|----------------------|----------|-----------------|-------|
| | | | MS | EMVNS | MS | EMVNS | MS | EMVNS |
| 4G2MM | 200 | 1 | -743.1162 | -743.1162 | 75.1187 | 50.8358 | 39 | 92 |
| | | 2 | -749.2199 | -749.2199 | 86.3220 | 64.6878 | 28 | 92 |
| | | 3 | -756.3678 | -756.3678 | 79.3128 | 68.4991 | 29 | 64 |
| | | 4 | -760.3703 | -760.3703 | 91.1277 | 55.7623 | 38 | 81 |
| | | 5 | -763.3643 | -763.3643 | 62.8596 | 58.2318 | 16 | 91 |
| | | 6 | -741.3652 | -741.3652 | 79.6966 | 20.2514 | 24 | 91 |
| | | 7 | -752.9941 | -752.9941 | 81.2247 | 34.6621 | 31 | 82 |
| | | 8 | -752.2411 | -752.2411 | 74.3248 | 56.7752 | 30 | 74 |
| | | 9 | -748.2610 | -748.2610 | 65.2544 | 12.5411 | 34 | 93 |
| | | 10 | -758.6311 | -758.6311 | 80.9817 | 70.2286 | 31 | 87 |
| 6G2MM | 500 | 1 | -1.0277e+03 | -1.0277e+03 | 403.7125 | 4.5753 | 1 | 98 |
| | | 2 | -977.9569 | -981.0494 | 405.4793 | 4.3618 | 3 | 86 |
| | | 3 | -1.0408e+03 | -1.0400e+03 | 399.9315 | 126.5601 | 1 | 5 |
| | | 4 | -975.4794 | -970.2479 | 429.9173 | 7.1039 | 1 | 12 |
| | | 5 | -998.3713 | -995.6342 | 423.5845 | 243.0707 | 1 | 41 |
| | | 6 | -989.1698 | -986.5160 | 417.2788 | 12.1357 | 1 | 68 |
| | | 7 | -995.4418 | -994.0682 | 434.6577 | 191.5254 | 1 | 31 |
| | | 8 | -1.0046e+03 | -995.0639 | 446.2635 | 416.3952 | 1 | 31 |
| | | 9 | -1.0188e+03 | -1.0178e+03 | 429.0988 | 13.4971 | 1 | 48 |
| | | 10 | -1.0463e+03 | -1.0451e+03 | 401.8871 | 4.0569 | 1 | 85 |
| 8G2MM | 700 | 1 | -958.6579 | -958.6579 | 1.0557e+03 | 75.3492 | 2 | 54 |
| | | 2 | -1.0081e+03 | -1.0081e+03 | 967.5115 | 73.8963 | 10 | 86 |
| | | 3 | -1.0472e+03 | -1.0472e+03 | 954.5972 | 125.2951 | 5 | 89 |
| | | 4 | -1.0077e+03 | -1.0077e+03 | 972.6653 | 147.8416 | 6 | 95 |
| | | 5 | -1.0040e+03 | -1.0040e+03 | 986.8095 | 201.4950 | 11 | 49 |
| | | 6 | -1.0090e+03 | -1.0090e+03 | 983.9204 | 881.9820 | 5 | 75 |
| | | 7 | -1.0084e+03 | -1.0732e+03 | 994.5039 | 10.3180 | 6 | 63 |
| | | 8 | -984.7680 | -984.7680 | 948.6578 | 76.3896 | 7 | 61 |
| | | 9 | -1.0171e+03 | -1.0171e+03 | 990.3082 | 744.9327 | 7 | 48 |
| | | 10 | -978.5214 | -978.5214 | 985.9514 | 601.1251 | 11 | 64 |
| 10G2MM | 800 | 1 | -1.1093e+03 | -1.1065e+03 | 1.2574e+03 | 11.4995 | 1 | 24 |
| | | 2 | -1.1698e+03 | -1.1685e+03 | 1.2375e+03 | 6.0270 | 1 | 43 |
| | | 3 | -1.1093e+03 | -1.1076e+03 | 1.2574e+03 | 166.1510 | 1 | 19 |
| | | 4 | -1.1848e+03 | -1.1711e+03 | 1.2309e+03 | 18.0059 | 1 | 84 |
| | | 5 | -1.1469e+03 | -1.1466e+03 | 1.2061e+03 | 54.1790 | 1 | 8 |
| | | 6 | -1.1682e+03 | -1.1784e+03 | 1.2315e+03 | 971.8920 | 1 | 16 |
| | | 7 | -1.1405e+03 | -1.1329e+03 | 1.2375e+03 | 31.4003 | 1 | 15 |
| | | 8 | -1.1202e+03 | -1.1256e+03 | 1.2888e+03 | 30.4336 | 1 | 22 |
| | | 9 | -1.1308e+03 | -1.1074e+03 | 1.2363e+03 | 33.1973 | 1 | 28 |
| | | 10 | -1.0882e+03 | -1.0968e+03 | 1.3139e+03 | 5.9767 | 1 | 33 |

References

- [1] Aitkin, M., Anderson, D., and Hinde, J.: Statistical Modelling of Data on Teaching Styles (with discussion), *J. Roy. Stat.Soc, A* 144, 419–461 (1981).
- [2] Aitkin, M., and Rubin, D.B.: Estimation and Hypothesis Testing in Finite Mixture Distributions, *Journal of the Royal Statistical Society, Series B*, 47, 67–75 (1985).
- [3] Basford, K.E., and McLachlan, G.J.: The Mixture Method of Clustering Applied to Three-Way Data, *Journal of Classification*, 2, 109–125 (1985).
- [4] Berchtold A.: Optimisation of mixture models: Comparison of different strategies, *Computational Statistics*, 19, 385–406 (2004).
- [5] Biernacki, C., Celeux, G., Govaert, G.: Choosing starting values for the EM algorithm for getting the highest likelihood in multivariate Gaussian mixture models, *Comput. Statist. Data Anal.*, 41, 561–575 (2003).
- [6] Biernacki, C.: Initializing em using the properties of its trajectories in gaussian mixtures, *Statistics and Computing*, 14(3), 267–279 (2004).
- [7] Bhattacharya, C.G.: A Simple Method for Resolution of a Distribution into its Gaussian Components, *Biometrics*, 23, 115–135 (1967).
- [8] Blekas, K., and Lagaris, I.E.: Split-Merge Incremental Learning (SMILE) of Mixture Models, *Inter. Conference on Artificial Neural Networks (ICANN)*, Porto 2007, *Lecture Notes on Artificial Neural Networks*, 4669, 291–300 (2007).
- [9] Boese, K.D., Kahng, A.B., and Muddu, S.: A New Adaptive Multi-Start Technique for Combinatorial Global Optimizations, *Oper. Res. Lett.*, 16, 101–113 (1994).
- [10] Böhning D.: *Computer-Assisted Analysis of Mixtures and Applications: Meta-Analysis, Disease Mapping and Others*. Chapman and Hall/CRC, New York (1999).
- [11] Cassie, R.M.: Some Uses of Probability Paper for the Graphical Analysis of Polymodel Frequency Distributions, *Australian Journal of Marine and Freshwater Research*, 5, 513–522, (1954).
- [12] Cohen, A.C. Jr.: Estimation in mixtures of two normal distributions, *Technometrics*, 9, 15–28 (1967).
- [13] Day, N.E.: Estimating the Components of a Mixture of two Normal Distributions, *Biometrika*, 56, 463–474 (1969).
- [14] Dempster, A.P., Laird, N.M., and Rubin, D.B.: Maximum likelihood from incomplete data via the em algorithm, *Journal of the Royal Statistical Society, Series B*, 39(1), 1, November (1977).
- [15] Do, K.-A., McLachlan G.J., Bean, R., and Wen, S.: Application of Gene Shaving and Mixture Models to Cluster Microarray Gene Expression Data, *Cancer Informatics*, 2, 25–43 (2007).
- [16] Everit, B.S., and Hand, D.J.: *Finite Mixture Distributions*, London: Chapman and Hall (1981).
- [17] Everit, B.S.: Maximum Likelihood Estimation of the Parameters in a Mixture of two Univariate Normal Distributions: A Comparison of Different Algorithms, *Statistician*, 33, 205–215 (1984).
- [18] Feo, T., and Resende M.G.C.: Greedy Randomized Adaptive Search Procedures, *Journal of Global Optimization*, 2, 1–27 (1995).
- [19] Finch, S., Mendell, N., and Thode, H.: Probabilistic measures of adequacy of a numerical search for a global maximum, *J. Amer. Statist. Assoc.*, 84, 1020–1023 (1989).
- [20] Fisher, R.A.: The Case of Zero Survivors, (Appendix to Bliss, C.I. (1935)), *Annals of Applied Biology*, 22, 164–165 (1935).
- [21] Fowlkes, E.B.: Some Methods for Studying Mixtures of two Normal (Lognormal) Distributions, *Journal of the American Statistical Association*, 74, 561–575 (1979).
- [22] Fryer, I., and Robertson, C.A.: A Comparison of Some Methods for Estimating Mixed Normal Distributions, *Biometrika*, 59, 639–648 (1972).
- [23] Glover, F.: Heuristics for Integer Programming Using Surrogate Constraints, *Decision Sciences*, 8, 156–166 (1977).

- [24] Glover, F.: Tabu Search: A tutorial, *Interfaces*, 20, 74–94 (1990).
- [25] Glover, F.: Multi-Start and strategic oscillation methods. Principles to exploit adaptive memory. *Computing tools for modeling optimization and simulation*, Eds. Laguna and Gonzalez-Velarde, Kluwer, 1–25 (2000).
- [26] Goodman, L.A.: Exploratory Latent Structure Analysis Using Both Identifiable and Unidentifiable Models, *Biometrika*, 61, 215–231 (1974).
- [27] Hansen, P., and Mladenović, N.: An Introduction to Variable Neighborhood Search. In S. Voss et al. (eds.), *Metaheuristics, Advances and Trends in Local Search Paradigms for Optimization*. Dordrecht: Kluwer, 433–458 (1998).
- [28] Hansen, P., and Mladenović, N.: Variable neighborhood search: Principles and applications. *European Journal of Operational Research*, 130, 449–467 (2001).
- [29] Hansen, P., and Mladenović, N.: Tutorial on variable neighborhood search. Technical Report G-2003-46, *Les Cahiers du GERAD* (2003).
- [30] Hansen, P., Mladenović, N., and Moreno Pérez, J.: Variable neighborhood search. *European Journal of Operational Research*, 191, 593–595 (2008).
- [31] Harding, I.P.: The Use of Probability Paper for the Graphical Analysis of Polymodel Frequency Distributions, *Journal of the Marine Biological Association (UK)*, 28, 141–153 (1948).
- [32] Hartley, H.: Maximum likelihood estimates from incomplete data. *Biometrics*, 14, 174–94 (1958).
- [33] Hasselblad, V.: Estimation of Parameters for a Mixture of Normal Distributions, *Technometrics*, 8, 431–444 (1966).
- [34] Hasselblad, V.: Estimation of Finite Mixtures of Distributions from the Exponential Family, *Journal of the American Statistical Association*, 64, 1459–1471 (1969).
- [35] Jank, W.S.: Ascent EM for fast and global model-based clustering: An application to curve-clustering of online auctions, *Computational Statistics and Data Analysis*, 51 747–761 (2006).
- [36] Karlis, D., and Xekalaki, E.: Choosing initial values for the EM algorithm for finite mixtures. *Comput. Statist. Data Anal.*, 41, 577 (2003).
- [37] Kirkpatrick, S., Gelatt, C.D., and Vecchi, M.P.: Optimization by Simulated Annealing, *Science*, 220, 671–680 (1983).
- [38] Lazarsfeld, P.F., and Henry, N.W.: *Latent Structure Analysis*, Boston, Houghton-Mifflin (1968).
- [39] Little, R.J.A., and Rubin, D.B.: *Statistical Analysis with Missing Data*. Second Edition. New York, John Wiley and Sons, Inc. (2002).
- [40] Hosmer, D.W. Jr.: On MLE of the parameters of a mixture of two normal distributions when the sample size is small, *Communications in Statistics—Theory and Methods*, 1, 217–227 (1973).
- [41] McKendrick, A.G.: Applications of Mathematics to Medical Problems, *Proceedings of the Erlinburgh Mathematical Society*, 44, 98–130 (1926).
- [42] McLachlan, G.J., and Krishnam, T.: *The EM algorithm and Extensions*. Wiley, New York (1997).
- [43] McLachlan, G.J.: On the choice of initial values for the EM algorithm in fitting mixture models, *The Statistician* 37, 417–425 (1988).
- [44] McLachlan, G.J., and Basford, K.E.: *Mixture Models: Inference and Applications to Clustering*. Marcel Dekker, New York (1988).
- [45] McLachlan, G., and Peel, D.: *Finite Mixture Models*, New-York, Wiley (2000).
- [46] McLachlan, G.J., Ng, S.K., and Bean, R.: Robust Cluster Analysis via Mixture Models, *Austrian Journal of Statistics*, 35(2–3), 157–3174 (2006).
- [47] Mendenhall, W., and Hadere, R.J.: Estimation of parameters of mixed exponentially distributed failure time distributions from censored life test data, *Biometrika*, 45, 504–520 (1958).
- [48] Meng, X.L., and Rubin, D.: Maximum likelihood estimation via the ECM algorithm: a general framework, *Biometrika*, 80, 267–278 (1993).

- [49] Meng, X. L., and van Dyk, D.: The EM algorithm – An old folk song sung to a fast new tune (with discussion), *Journal of the Royal Statistical Society B*, 59, 511–567 (1997).
- [50] Meng X.L. and Pedlow S.: A bibliographic review with missing articles, *Proc. Statist. Comput.Sect. Am. Statist. Ass.*, 39, 24–27 (1992).
- [51] Meng, X.L.: Missing data: Dial M for ???, *Journal of the American Statistical Association*, 95, 1325–1330 (2000).
- [52] Mladenović, N., and Hansen, P.: Variable neighborhood search. *Computers and Operations Research* 24, 1097–1100 (1997).
- [53] Newcomb, S.: A Generalized Theory of the Combination of Observations So As To Obtain the Best Result, *American Journal of Mathematics*, 8, 343–366 (1886).
- [54] Pearson, K.: Contributions to the Mathematical Theory of Evolution, *Philosophical Transactions, A*, 185, 71–110 (1894).
- [55] Peel, D., and McLachlan, G.J.: Robust mixture modeling using the t distribution, *Statistical Computing*, 10, 335–344 (2000).
- [56] Rao, C.R.: The utilization of multiple measurements in problems of biological classification, *J. Royal Statist. Soc. Ser. B*, 10, 159–193 (1948).
- [57] Redner, R.A., and Walker, H.F.: Mixture densities, maximum likelihood and the EM algorithm, *SIAM Review*, 26, 195–239 (1984).
- [58] Rinnooy Kan, A.H.G., and Timmer, G.T.: *Global Optimization Handbooks in operations research and management science*, Vol. 1, Ed. Rinnooy Kan and Todds, North Holland, pp. 631–662 (1989).
- [59] Stigler, S.M.: Studies in the History of Probability and Statistics. XXXII Laplace, Fisher, and the discover of the concept of sufficiency, *Biometrika*, 60, 439–445 (1973).
- [60] Tanand, W.Y., and Change, W.C.: Convolution approach to genetic analysis of quantitative characters of self-fertilized population, *Biometrics*, 28, 1073–1090 (1972).
- [61] Thomas, E.A.C.: Mathematical Models for the Clustered Firing of Single Cortical Neurons, *British Journal of Mathematical and Statistical Psychology*, 19, 151–162 (1966).
- [62] Titterington, D., Smith, A., and Makov, U.: *Statistical Analysis of Finite Distributions*, New York, NY, John Wiley (1985).
- [63] Titterington, D.M.: Some Recent Research in the Analysis of Mixture Distributions, *Statistics*, 4, 619–641 (1990).
- [64] Thomas, E.A.C.: Mathematical Models for the Clustered Firing of Single Cortical Neurons, *British Journal of Mathematical and Statistical Psychology*, 19, 151–162 (1966).
- [65] Teicher, H.: Identifiability of Mixtures, *Annals of Mathematical Statistics*, 31, 55–73 (1961).
- [66] Ueda, N., and Nakano, R.: Deterministic annealing EM algorithm, *Neural Networks*, 11, 271–282 (1998).
- [67] Ueda, N., Nakano, R., Ghahramani, Z., and Hinton, G.E.: SMEM algorithm for mixture models, *Neural Computation*, 12, 2109–2128 (2000).
- [68] Vlassis, N., Likas, A.: A greedy EM algorithm for Gaussian mixture learning, *Neural Process. Lett.*, 15, 77–87 (2002).
- [69] Wolfe, J.H.: Pattern Clustering by Multivariate Mixture Analysis, *Multivariate Behavioral Research*, 5, 329–350 (1970).
- [70] Woodruff, D.L., Reiners, T.: Experiments with, and on, algorithms for maximum likelihood clustering, *Comput. Statist. Data Analysis*, 47, 237–253 (2004).
- [71] Wright, K., and Kennedy, W.J.: An interval analysis approach to the EM algorithm, *Journal of Computational and Graphical Statistics*, 9, 303–318 (2000).
- [72] Wu, C.F.J.: On the convergence properties of the EM algorithm, *Annals of Statistics*, 11, 95–103 (1983).
- [73] Xu, L., and Jordan, M.I.: On convergence properties of the em algorithm for gaussian mixtures, *Neural Computation*, 8, 129–151 (1996).